

Clinical Research Article

The Effects of Vitamin D Supplementation and 25-Hydroxyvitamin D Levels on the Risk of Myocardial Infarction and Mortality

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Abbreviations: (25-OH)D, 25-hydroxyvitamin D; BMI, body mass index; CDW, Corporate Data Warehouse; HR, hazard ratio; ICD, International Classification of Disease; IPTW, inverse probability of treatment weight; KM, Kaplan–Meier; LDL, low-density lipoprotein; MI, myocardial infarction; VDT, vitamin D treatment; VHA, Veterans Health Administration; VIDA, Effect of monthly high-dose vitamin D supplementation on cardiovascular disease in the Vitamin D Assessment Study; VINCI, Veterans Administration Informatics and Computing Infrastructure; VITAL, Vitamin D supplements and prevention of cancer and cardiovascular disease; Vit-D, vitamin D.

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Abstract

Objective: The aim of the study was to examine the effects of the vitamin D (Vit-D) treatment and nontreatment on Vit-D–deficient patients without a prior history of myocardial infarction (MI).

Materials and Methods: This was a retrospective, observational, nested case–control study of patients (N = 20 025) with low 25-hydroxyvitamin D ([25-OH]D) levels (<20 ng/mL) who received care at the Veterans Health Administration from 1999 to 2018. Patients were divided into 3 groups: Group A (untreated, levels ≤20 ng/mL), Group B (treated, levels 21–29 ng/mL), and Group C (treated, levels ≥30 ng/mL). The risk of MI and all-cause mortality were compared utilizing propensity score–weighted Cox proportional hazard models.

Results. Among the cohort of 20 025 patients, the risk of MI was significantly lower in Group C than in Group B (hazard ratio [HR] 0.65, 95% CI 0.49–0.85, *P* = .002) and Group A (HR 0.73, 95% CI 0.55–0.96), *P* = .02). There was no difference in the risk of MI between Group B and Group A (HR 1.14, 95% CI 0.91–1.42, *P* = 0.24). Compared with Group A, both Group B (HR 0.59, 95% CI 0.54–0.63, *P* < .001) and Group C (HR 0.61, 95% CI 0.56–0.67, *P* < .001) had significantly lower all-cause mortality. There was no difference in all-cause mortality between Group B and Group C (HR 0.99, 95% CI 0.89–1.09, *P* = .78).

Conclusions. In patients with Vit-D deficiency and no prior history of MI, treatment to the (25-OH)D level of >20 ng/mL and >30 ng/mL was associated with a significantly lower risk of all-cause mortality. The lower risk of MI was observed only in individuals maintaining (25-OH)D levels \geq 30 ng/mL.

Key Words: Vitamin D, myocardial infarction, all-cause mortality, primary prevention, cardiovascular disease

There is substantial evidence implicating vitamin D (Vit-D) levels in the pathogenesis of cardiovascular risk factors such as diabetes, hypertension, hyperlipidemia, chronic kidney disease, and obesity [1, 2]. Furthermore, experimental studies suggest that Vit-D may participate in pathways associated with atherosclerosis by influencing cellular growth, oxidative stress, membrane transport, cell adhesion, and gene regulation. Direct effects of Vit-D on cardiomyocytes and vascular endothelial cells were reported via Vit-D receptors [1]. The Endocrine Society defines 25-hydroxyvitamin D ((25-OH)D) levels \leq 20 ng/mL as deficiency, levels 21–29 ng/mL as insufficiency, and levels \geq 30 ng/mL as optimal [3]. However, data regarding the association of the (25-OH)D levels and Vit-D supplementation with myocardial infarction (MI) and mortality remains controversial [4–14].

Several meta-analyses of epidemiological studies suggested that Vit-D deficiency is associated with an increased risk of MI and cardiovascular mortality [8, 12, 14]. One meta-analysis suggested that there is generally a linear, inverse association between circulating (25-OH)D levels and the risk of cardiovascular disease [13]. A Cochrane meta-analysis showed that Vit-D treatment significantly reduced mortality in subgroups of patients with a pretreatment level below 20 ng/mL [7]. However, in several randomized controlled trials, supplementation of Vit-D did not result in lower cardiovascular events and mortality [4–6]. It is important to note that majority of these randomized controlled trials had included patients who already had optimal baseline (25-OH)D levels, with most patients in these trials having pretreatment (25-OH)D levels above 25 to 30 ng/mL [4, 15]. Additionally, in the majority of these clinical trials, post-treatment follow-up of (25-OH)D was not measured to account for effective supplementation and had a short-term follow-up [4]. Even in the VITAL (vit-D and omega-3 trial) and the VIDA (vit-D assessment) trials, only a small subset of the study population (6.3% and 8.6%, respectively) had a repeat measurement of post-treatment (25-OH)D level performed [5, 6]. It is also worth noting that in several studies, the association between Vit-D and the risk of MI was apparent only after long-term follow-up [8, 10, 11]. Additionally, there are limited data available comparing the outcome of MI and mortality with respect to the levels of (25-OH)D achieved and maintained after Vit-D supplementation.

To address this gap in knowledge, we conducted a large retrospective analysis with long-term follow-up in patients

with low baseline Vit-D level who had at least 2 separate measurements of (25-OH)D levels to confirm their status and to measure the effect of Vit-D supplementation on (25-OH)D levels. The goal of our study was to examine the effects of Vit-D treatment (VDT) and lack of VDT on all-cause mortality and MI in Vit-D-deficient patients without prior history of MI in relation to 3 different reference levels of (25-OH)D as defined by the Endocrine Society.

Material and Methods

In this retrospective, observational, nested case-control study we leveraged clinical data ascertained from the Veterans Health Administration (VHA) Corporate Data Warehouse (CDW) through the Veterans Administration Informatics and Computing Infrastructure (VINCI) [16]. VINCI hosts the data, facilitates analysis while ensuring the privacy of veterans as well as data security [16]. This study was approved by the Institutional Review Board of the Kansas City Veterans Affairs Medical Center, MO, USA.

Study Design

This study was designed to examine the association of Vit-D with MI and all-cause mortality among different subpopulations of treated and untreated patients. The incidence of MI and coexisting conditions were based on the International Classification of Disease 9th and 10th revision (ICD-9 and ICD-10) codes. All the study population had (25-OH)D levels checked on at least 2 separate occasions to be included in the study.

Study Population

Inclusion criteria

Our study included veterans (1) who received their medical care at the VHA between December 1999 and December 2018, (2) who were tested for (25-OH)D levels, (3) those whose baseline level of (25-OH)D was \leq 20 ng/mL, and (4) whose age was >18 years.

Exclusion criteria

We excluded (1) patients on VDT prior to the index (25-OH)D level, (2) those who had MI before the first study

date, and (3) those who did not have follow-up 25(OH)D testing done after initiation of treatment. We also excluded patients who had a baseline or follow-up 25(OH)D level ≥ 100 ng/mL. Although the safe upper level of (25-OH)D for avoiding hypercalcemia is uncertain, Vit-D intoxication is usually observed in (25-OH)D above 150 ng/mL [17]. Hence, an upper limit of 100 ng/mL has been suggested to provide a safety margin in reducing effects of hypercalcemia [3].

The study population was divided into 3 groups: (1) patients who did not receive VDT and their (25-OH)D levels remained ≤ 20 ng/mL on follow-up (untreated, level ≤ 20) (Group A), (2) patients who received VDT and their (25-OH)D levels remained between 21 and 29 ng/mL upon follow-up (treated, level 21-29) (Group B), and (3) patients who received VDT and their (25-OH)D level remained ≥ 30 ng/mL upon follow-up (treated, level ≥ 30) (Group C). (Fig. 1)

Rationale for the 25-Hydroxyvitamin D Cut-off Levels

In the general population, the Endocrine Society clinical practice guidelines define Vit-D deficiency as the (25-OH)

D levels ≤ 20 ng/mL, levels 21 to 29 ng/mL as insufficiency, and levels ≥ 30 ng/mL as optimal [3]. We utilized these definitions in our study, and patients with baseline (25-OH)D level ≤ 20 ng/mL were defined to have low Vit-D.

Ascertainment of Vitamin D Treatment Exposure

The use of Vit-D supplementation was ascertained from the medication prescription of patient medical records. Any patient who received cholecalciferol or ergocalciferol (capsule or tablet), defined by release of the medication by the pharmacy, was considered to have been treated.

Outcome Measures

The outcomes of the study were all-cause mortality and MI across Vit-D levels. Deaths from any cause were determined using dates of death in CDW data augmented with vital status files. MI was identified using the International Classification of Diseases, Ninth and Tenth Revisions, Clinical Modification (ICD-9-CM and ICD-10-CM) code system. All serum (25-OH)D levels were measured at the VA healthcare clinical laboratories utilizing uniform standardized techniques.

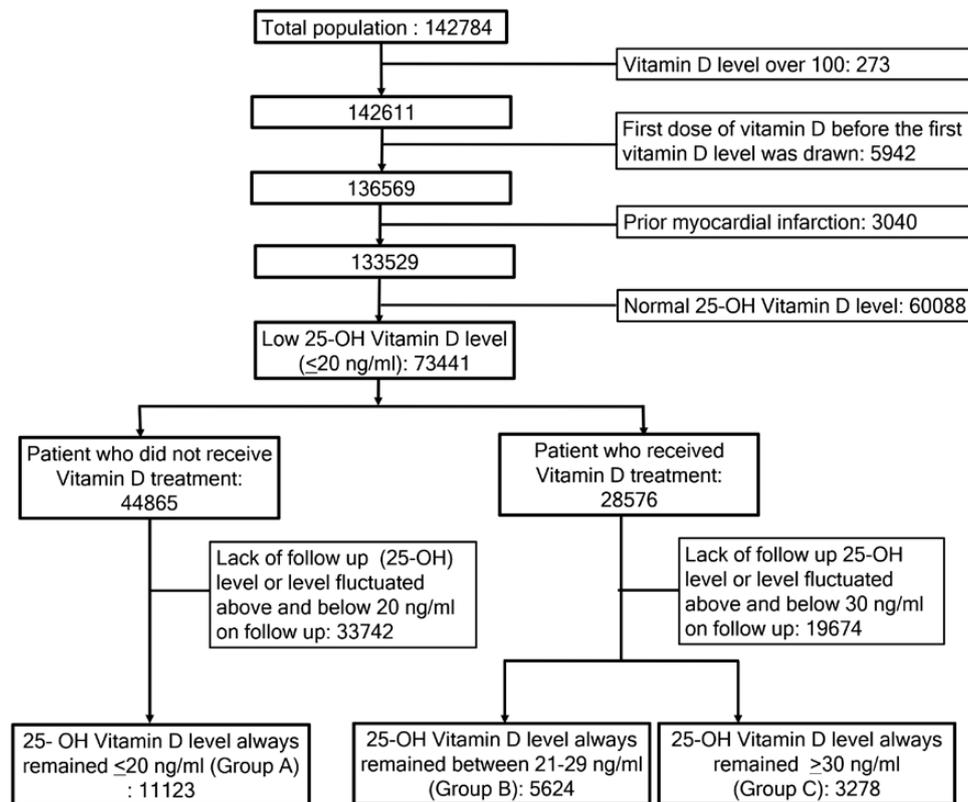


Figure 1. Methodology and patients selection process. Selection of the study population. We excluded 123 692 patients who met various exclusion criteria. A total of 19 092 patients met our inclusion criteria and were divided into 3 subgroups according to their treatment status and follow-up (25-OH)D levels: Group A, untreated, level < 20 ng/mL; Group B, treated, level 20-29 ng/mL; and Group C, treated, level ≥ 30 ng/mL.

Statistical Analysis

All categorical and continuous variables were reported as percentages and mean with SD, respectively. Differences in mean and percentage were assessed using the Student's t-test and Pearson chi-squared test. Univariate and multivariable Cox proportional hazard regression models were utilized to assess the differences between the tested groups. Propensity scores were used to correct for potential systematic differences between the comparison groups. The patient's propensity scores for receiving VDT were computed and adjusted for the covariates in a logistic regression analysis. The covariates included were age, sex, body mass index (BMI), hypertension, diabetes, coronary artery disease, congestive heart failure, peripheral arterial disease, chronic kidney disease, chronic obstructive pulmonary disease, smoking, concomitant therapies (aspirin, statin, and beta-blockers) and the low-density lipoprotein (LDL) cholesterol levels.

Additionally, we utilized propensity score-weighted, stabilized inverse probability of treatment weights (IPTWs) to obtain unbiased estimates of the treatment effects [18, 19]. IPTW accounts for confounding, distributing the weights according to the sample representation in which the treatment is independent of the measured confounders. The use of stabilized IPTW helped control for the imbalances between the comparison groups as shown in Table 1. Stabilized IPTW was also applied to the survival analysis to derive Kaplan–Meier (KM) survival curves comparing event-free survival time. STATA 15 (Stata Corp, College Station, TX) was used for statistical analysis. A 2-sided P value of < 0.05 was chosen as the level of statistical significance.

Results

Description of the Patient Cohort

Figure 1 outlines the study subject enrollment details. A total of 142 784 patients had (25-OH)D levels tested. There were 273 patients with baseline or follow-up (25-OH)D level above 100 ng/mL who were excluded to avoid the effect of VDT toxicity. We then excluded 5942 patients whose pretreatment baseline (25-OH)D level could not be ascertained. Subsequently, we excluded 3040 patients who had prior MI as our study was focused on role of Vit-D in primary prevention. Among the remaining 133 529 patients, 44.99% (60 088) patients had normal (25-OH)D at baseline and were excluded. There were 73 441 patients with low (25-OH)D at baseline, among those 38.91% (28 576) patients received VDT and 61.09% (44 865) patients were identified as naïve. From these 2 groups of patients, 54 349 patients

were excluded (1) as they did not have a follow-up (25-OH)D level measured, (2) a follow-up (25-OH)D level fluctuated above the prespecified threshold for the group, or (3) the timestamp for the follow-up (25-OH)D level was missing. The remaining 19 092 patients were categorized into 3 study groups. There were 11 119 who did not receive treatment and had (25-OH)D levels which remained below 20 ng/mL (Group A). Among the patients who received VDT, 5623 patients had follow-up (25-OH)D level that remained between 20 and 29 ng/mL (Group B) and 3277 patients had a follow-up (25-OH)D level that remained > 30 ng/mL (Group C). The mean time between the diagnosis of low 25(OH)D level and follow-up (25-OH)D level was 2.14 years (SD 2.06). The mean numbers of times (25-OH)D levels were repeated after the diagnosis of low baseline (25-OH)D level was 2.62 (SD 2.48). The median time between the last sample collection for (25-OH)D level and MI was 1.17 (25-75% 0.42-2.52) years.

Baseline Characteristics of the Patients

The baseline characteristics of the 3 comparison groups are shown in Table 1. Utilizing the stabilized IPTW, we balanced and matched for differences in age, sex, BMI, patients' comorbidities, concomitant therapies with aspirin, statin, and beta blockers as well as LDL level by ensuring that the cohorts were well matched ($P > .05$). (Table 1)

Association of (25-OH)D levels With Myocardial Infarction

The risk of myocardial infarction in the study groups is presented in Table 2. The risk of MI in Group B (treated, level 20-29 ng/mL) was not different from that of Group A (untreated, level < 20 ng/mL) (HR 1.14, 95% CI 0.91-1.42, $P = .24$). However, Group C (treated, level > 30 ng/mL) had a lower risk of MI than both Group B (treated, level 20-29 ng/mL) (HR 0.65, 95% CI 0.49-0.85, $P = .002$) as well as Group A (untreated, level < 20 ng/mL) (HR 0.73, 95% CI 0.55-0.96, $P = .02$). A comparison of the probability of MI-free survival with a KM curve among the 3 groups is shown in Fig. 2. KM curves show that Group C (treated, level > 30 ng/mL group) had significantly higher MI free survival than Group B (treated, level 20-29 ng/mL) (log-rank $P < .001$) and Group A (untreated, level < 20 ng/mL) (log-rank $P = .03$) group. There was no significant difference in MI-free survival between Group B (treated, level 20-29 ng/mL) and Group A (untreated, level < 20 ng/mL) (log-rank $P = .10$).

Table 1. Baseline characteristics of all patients in the study unadjusted and stabilized inverse probability of treatment weight adjusted

| Patient characteristics | Unmatched cohort | | | Propensity-matched cohort (stabilized IPTW) | | |
|--|--------------------------------|--------------------------------|---------|---|--------------------------------|---------|
| | Untreated, level ≤20 (Group A) | Treated, level 21-29 (Group B) | P value | Untreated, level ≤20 (Group A) | Treated, level 21-29 (Group B) | P value |
| Untreated, level ≤20 ng/mL vs treated, level 21-29 ng/mL (Group A vs Group B) | | | | | | |
| Number of patients (N) | 11 119 | 5623 | | 10 064 | 5067 | |
| Age ≥ 50 years, n (%) | 7946 (71.5) | 4000 (71.1) | .67 | 7256 (72.1) | 3663 (72.3) | .73 |
| Age, mean years (SD) | 57.7 (17.9) | 56.9 (16.3) | | 57.6 (17.5) | 58.0 (16.5) | |
| Male, n (%) | 8537 (76.8) | 3715 (66.1) | <.001 | 7367 (73.2) | 3694 (72.9) | .75 |
| Body mass index, kg/m ² , mean (SD) | 29.9 (6.7) | 31.3 (6.7) | <.001 | 30.5 (6.8) | 30.6 (6.6) | .61 |
| BMI ≥30 | 5560 (50) | 3138 (55.8) | <.001 | 4992 (49.6) | 2513 (49.6) | .97 |
| Follow-up time (years), mean (SD) | 3.6 (2.9) | 4.6 (2.8) | | 3.6 (3) | 4.5 (2.8) | |
| Comorbidities | | | | | | |
| Hypertension, n (%) | 6266 (56.4) | 3299 (58.7) | .004 | 6089 (60.5) | 3071 (60.6) | .88 |
| Diabetes mellitus, n (%) | 3201 (28.8) | 1712 (30.5) | .03 | 3150 (31.3) | 1586 (31.3) | .97 |
| Coronary artery disease, n (%) | 1669 (15.0) | 816 (14.5) | .39 | 1600 (15.9) | 801 (15.8) | .92 |
| Congestive heart failure, n (%) | 807 (7.3) | 338 (6.0) | .003 | 725 (7.2) | 365 (7.2) | .88 |
| Peripheral vascular disease, n (%) | 889 (8.0) | 425 (7.6) | .32 | 835 (8.3) | 421 (8.3) | .93 |
| Chronic kidney disease, n (%) | 1097 (9.9) | 479 (8.5) | .005 | 976 (9.7) | 502 (9.9) | .79 |
| Chronic obstructive pulmonary disease, n (%) | 495 (4.5) | 201 (3.6) | .007 | 453 (4.5) | 228 (4.5) | .97 |
| Smoking, n (%) | 2675 (24.1) | 1447 (25.7) | .02 | 2647 (26.3) | 1338 (26.4) | .90 |
| Concomitant therapy | | | | | | |
| Aspirin, n (%) | 6776 (60.9) | 4164 (74.1) | <.001 | 6783 (67.4) | 3435 (67.8) | .63 |
| Statin, n (%) | 5996 (53.9) | 3697 (65.8) | .001 | 6059 (60.1) | 3060 (60.4) | .81 |
| Beta blockers, n (%) | 4475 (40.3) | 2606 (46.4) | <.001 | 4398 (43.7) | 2224 (43.9) | .73 |
| Laboratory findings | | | | | | |
| Low density lipoprotein, mg/dl, mean (SD) | 107.3 (36.5) | 109.4 (36.5) | <.001 | 107.7 (36.4) | 107.6 (36.3) | .91 |
| Treated, level 21-29 ng/mL vs Treated, level ≥30 ng/mL (Group B vs Group C) | | | | | | |
| | Treated, level 21-29 (Group B) | Treated, level ≥30 (Group C) | P value | Treated, level 21-29 (Group B) | Treated, level ≥30 (Group C) | P value |
| Number of patients (N) | 5623 | 3277 | | 5266 | 3088 | |
| Age ≥50 years, n (%) | 4000 (71.1) | 2703 (82.5) | <.001 | 3997 (75.9) | 2331 (75.5) | .66 |
| Age, mean years (SD) | 56.9 (16.3) | 62.2 (15.5) | | 58.9 (16.2) | 59.5 (15.8) | |
| Male, n (%) | 3715 (66.1) | 2272 (69.3) | .002 | 3565 (67.7) | 2084 (67.5) | .88 |
| BMI, kg/m ² , mean (SD) | 31.3 (6.7) | 29.3 (6.2) | <.001 | 30.6 (6.6) | 30.7 (6.8) | .61 |
| BMI ≥30 | 3138 (55.8) | 1432 (43.7) | <.001 | 2623 (49.9) | 1547 (50.1) | .9 |
| Follow-up time (years), mean (SD) | 4.6 (2.8) | 4.9 (3.1) | | 4.6 (2.8) | 4.9 (3.0) | |
| Comorbidities | | | | | | |
| Hypertension, n (%) | 3299 (58.7) | 2178 (66.5) | <.001 | 3365 (63.9) | 1967 (63.7) | .89 |
| Diabetes mellitus, n (%) | 1712 (30.5) | 1023 (31.2) | .45 | 1690 (32.1) | 997 (32.3) | .88 |
| Coronary artery disease, n (%) | 816 (14.5) | 621 (18.9) | <.001 | 885 (16.8) | 513 (16.6) | .79 |
| Congestive heart failure, n (%) | 338 (6.0) | 241 (7.4) | .01 | 3581 (6.8) | 207 (6.7) | .90 |
| Peripheral vascular disease, n (%) | 425 (7.6) | 258 (7.9) | .59 | 427 (8.1) | 244 (7.9) | .83 |
| Chronic kidney disease, n (%) | 479 (8.5) | 281 (8.6) | .93 | 469 (8.9) | 278 (9.0) | .87 |
| Chronic obstructive pulmonary disease, n (%) | 201 (3.6) | 144 (4.4) | .05 | 216 (4.1) | 127 (4.1) | .96 |
| Smoking, n (%) | 1447 (25.7) | 927 (28.3) | .009 | 1464 (27.8) | 853 (27.6) | .86 |
| Concomitant therapy | | | | | | |
| Aspirin, n (%) | 4164 (74.1) | 2367 (72.2) | .06 | 3913 (74.3) | 2297 (74.4) | .93 |
| Statin, n (%) | 3697 (65.8) | 2342 (71.5) | <.001 | 3639 (69.1) | 2125 (68.8) | .74 |
| Beta blockers, n (%) | 2606 (46.4) | 1609 (49.1) | .01 | 2538 (48.2) | 1482 (48.0) | .88 |
| Laboratory findings | | | | | | |
| Low density lipoprotein, mg/dl, mean (SD) | 109.4 (36.5) | 101.9 (35.3) | <.001 | 106.5 (36.0) | 106.7 (36.7) | .84 |

Table 1. Continued

Untreated, level ≤ 20 ng/mL vs Treated, level ≥ 30 ng/mL
(Group A vs Group C)

| | Untreated, level ≤ 20 (Group A) | Treated, level ≥ 30 (Group C) | <i>P</i> value | Untreated, level ≤ 20 (Group A) | Treated, level ≥ 30 (Group C) | <i>P</i> value |
|---|---|---------------------------------------|-------------------|---|---------------------------------------|-------------------|
| Number of patients (N) | 11 119 | 3277 | | 10 014 | 2942 | |
| Age ≥ 50 years, n (%) | 7946 (71.5) | 2703 (82.5) | <.001 | 7480 (74.7) | 2218 (75.4) | .53 |
| Age, mean years (SD) | 57.7 (17.9) | 62.2 (15.5) | | 58.8 (17.5) | 60.1 (16.2) | |
| Sex, n (%) | 8537 (76.8) | 2272 (69.3) | <.001 | 7540 (75.3) | 2198 (74.7) | .52 |
| Body mass index, kg/m ² , mean (SD) | 29.9 (6.7) | 29.3 (6.2) | <.001 | 29.9 (6.6) | 29.8 (6.5) | .72 |
| BMI ≥ 30 | 5560 (50) | 1432 (43.7) | <.001 | 4586 (45.8) | 1315 (44.7) | .31 |
| Follow-up time (years), mean (SD) | 3.6 (2.9) | 4.9 (3.1) | | 3.6 (3) | 4.7 (3.1) | |
| Comorbidities | | | | | | |
| Hypertension, n (%) | 6266 (56.4) | 2178 (66.5) | <.001 | 6239 (62.3) | 1853 (63.0) | .51 |
| Diabetes mellitus, n (%) | 3201 (28.8) | 1023 (31.2) | .007 | 3135 (31.3) | 936 (31.8) | .65 |
| Coronary artery disease, n (%) | 1669 (15) | 621 (18.9) | <.001 | 1702 (17) | 503 (17.1) | .95 |
| Congestive heart failure, n (%) | 807 (7.3) | 241 (7.4) | .87 | 771 (7.7) | 232 (7.9) | .79 |
| Peripheral vascular disease, n (%) | 889 (8) | 258 (7.9) | .82 | 841 (8.4) | 247 (8.4) | .99 |
| Chronic kidney disease, n (%) | 1097 (9.9) | 281 (8.6) | .03 | 991 (9.9) | 300 (10.2) | .59 |
| Chronic obstructive pulmonary disease, n (%) | 495 (4.5) | 144 (4.4) | .89 | 481 (4.8) | 153 (5.2) | .52 |
| Smoking, n (%) | 2675 (24.1) | 927 (28.3) | <.001 | 269 (26.9) | 809 (27.4) | .58 |
| Concomitant therapy | | | | | | |
| Aspirin, n (%) | 6776 (60.9) | 2367 (72.2) | <.001 | 6559 (65.5) | 1951 (66.3) | .46 |
| Statin, n (%) | 5996 (53.9) | 2342 (71.5) | <.001 | 6028 (60.2) | 1774 (60.3) | .99 |
| Beta blockers, n (%) | 4475 (40.2) | 1609 (49.1) | <.001 | 4366 (43.6) | 1294 (44.0) | .69 |
| Laboratory findings | | | | | | |
| Low density lipoprotein, mg/dL, mean (SD) | 107.3 (36.5) | 101.9 (35.3) | <.001 | 105.7 (36.1) | 105.6 (36.9) | .97 |

The covariates included were age, sex, body mass index (BMI), risk factors hypertension, diabetes mellitus, coronary artery disease, congestive heart failure, peripheral vascular disease, chronic kidney disease, chronic obstructive pulmonary disease, and smoking, concomitant therapies (aspirin, statin, and beta-blockers) and low-density lipoprotein.

Association of (25-OH)D Levels With and All-cause Mortality

Table 2 summarizes study group-stratified risk distribution for all-cause mortality. Compared with Group A (untreated, level < 20 ng/mL), the risk of all-cause mortality was significantly lower in both Group C (treated, level > 30 ng/mL) (HR 0.61, 95% CI 0.56-0.67, $P < .001$) and Group B (treated, level 20-29 ng/mL) (HR 0.59, 95% CI 0.54-0.63, $P < .001$). However, there was no significant difference in the risk of all-cause mortality between Group C (treated, level > 30 ng/mL) and Group B (treated, level 20-29 ng/mL) (HR 0.99, 95% CI 0.89-1.09, $P = .78$). Survival analysis with KM curve shows that the probability of survival was significantly higher in Group B (treated, level 20-29 ng/mL) (log-rank $P < .001$) and Group C (treated, level > 30 ng/mL) (log-rank $P < .001$) than Group A (untreated, level < 20 ng/mL). The probability of survival was not different between

Group B (treated, level 20-29 ng/mL and Group C (treated, level > 30 ng/mL) (log-rank $P = .78$). (Fig. 3)

Discussion

The current study examined the effects of nontreatment and treatment in Vit-D-deficient patients without a prior history of MI in relation to 3 different reference levels of (25-OH)D. Based on the long-term follow-up, our study found that the patients with post-treatment (25-OH)D levels at or above 30 ng/mL had lower incidence of MI and all-cause mortality. These results suggest that targeting 25(OH)D levels above 30 ng/mL might improve prognosis in the primary prevention setting among individuals with Vit-D deficiency.

There are conflicting data related to low Vit-D level and all-cause mortality [4-14]. Our study appears to unify and

Table 2. Hazard ratio for all-cause mortality and myocardial infarction among the propensity matched, stabilized inverse probability of treatment weighted subgroups

| Outcomes | All-cause mortality | | | Myocardial infarction | | |
|--|---------------------|-----------|---------|-----------------------|-----------|---------|
| | Hazard ratio | 95% CI | P value | Hazard ratio | 95% CI | P value |
| Comparing untreated, level ≤20 ng/mL (Group A) vs treated, level 21-29 ng/mL (Group B) (reference = untreated, level ≤20 ng/mL) | | | | | | |
| Propensity matched (stabilized IPTW) N = 10 064 vs 5067 | 0.59 | 0.54-0.63 | <.001 | 1.14 | 0.91-1.42 | .24 |
| Comparing treated, level 21-29 ng/mL (Group B) vs treated, level ≥30 ng/mL (Group C) (reference = treated, level 21-29 ng/mL) | | | | | | |
| Propensity matched (stabilized IPTW) N = 5266 vs. 3088 | 0.99 | 0.89-1.09 | .78 | 0.65 | 0.49-0.85 | .002 |
| Comparing untreated, level ≤20 ng/mL (Group A) vs treated, level ≥30 ng/mL (Group C) (reference = untreated, level ≤20 ng/mL) | | | | | | |
| Propensity matched (stabilized IPTW) N = 10 014 vs. 2942 | 0.61 | 0.56-0.67 | <.001 | 0.73 | 0.55-0.96 | .02 |

Abbreviation: IPTW, inverse probability of treatment weight.

provide an explanation for some of the contradictory data related to Vit-D and its association with MI and all-cause mortality. These data also provide a possible unique perspective regarding the association of MI and mortality in relation to Vit-D deficiency and Vit-D supplementation. Our finding of a significantly lower all-cause mortality when the Vit-D levels were maintained >20 ng/mL post-treatment compared with patients who were untreated and whose levels remained ≤20 ng/mL, is consistent with several prior prospective studies with long-term follow-up and meta-analyses of randomized studies [7, 9, 20, 21]. The Cochrane meta-analyses showed Vit-D supplementation in patients with (25-OH)D levels <20 ng/mL significantly lower all-cause mortality and this benefit was not seen in patients with (25-OH)D levels above 20 ng/mL [7]. In a 20-year-follow-up of the Third National Health and Nutrition Examination Survey (NHANES III) participants, (25-OH)D levels above 17.5 ng/mL were associated with lower all-cause mortality [22]. On the contrary, the VITamin D and Omega-3 Trial (VITAL) followed by the subsequent meta-analyses of clinical trials did not show any reduction in all-cause mortality with Vit-D supplementation [4, 5]. In these studies, however, all-cause mortality as an outcome was not specifically looked at in patients with baseline (25-OH)D levels ≤20 ng/mL [4, 20, 23]. The other reason for the difference could be that in these study population the baseline (25-OH)D levels were >20 ng/mL when VDT was initiated. For example, in the VITAL trial, the majority of patients had baseline (25-OH)D levels above 25-30 ng/mL, with only 12.7% of the population having (25-OH)D levels <20 ng/mL and 32% of patients between 20 and 29 ng/mL [5]. In our study, we found there was no difference in all-cause mortality among groups with

(25-OH)D levels between 21 and 29 ng/mL and >30 ng/mL after treatment. When the post-treatment (25-OH)D level >20 ng/mL was achieved, there was no added benefit on all-cause mortality with higher (25-OH)D levels. Therefore, it may be hypothesized that in Vit-D-deficient patients, a target (25-OH)D level of >20 ng/mL would be sufficient to obtain a mortality benefit.

We found a lower incidence of MI events in patients who had (25-OH)D levels at or above 30 ng/mL compared with those with the levels in the range 21 to 29 ng/mL as well as those with levels ≤20 ng/mL. These findings are consistent with prior studies demonstrating varying risk of MI upon long-term follow-up in patients according to their baseline (25-OH)D levels [8, 10, 11]. There was no difference in the risk of MI between patients with (25-OH)D levels maintained at ≤20 ng/mL and 21 to 29 ng/mL. Our findings suggest that (25-OH)D target level ≥30 ng/mL may provide protection against MI. Study by Brøndum-Jacobsen et al. also showed graded increase in MI and mortality with drop in (25-OH)D levels [8]. In our study, when compared with the ≤20 ng/mL group of patients, those with ≥30 ng/mL had both mortality and MI benefit. On the other hand, no significant difference in all-cause mortality was noted between the groups with (25-OH)D levels 20-29 ng/mL and ≥30 ng/mL, but there was a significant difference in the MI event rate, suggesting that the MI events between these 2 groups may not contribute significantly to the mortality. While observational studies demonstrate similar findings of lower MI risk with Vit-D supplementation [10, 11], this effect has not been replicated yet in randomized clinical trials [4, 5]. It is plausible that this discrepancy stems from the lack of standardized target levels for (25-OH)D across the studies. Further, risk

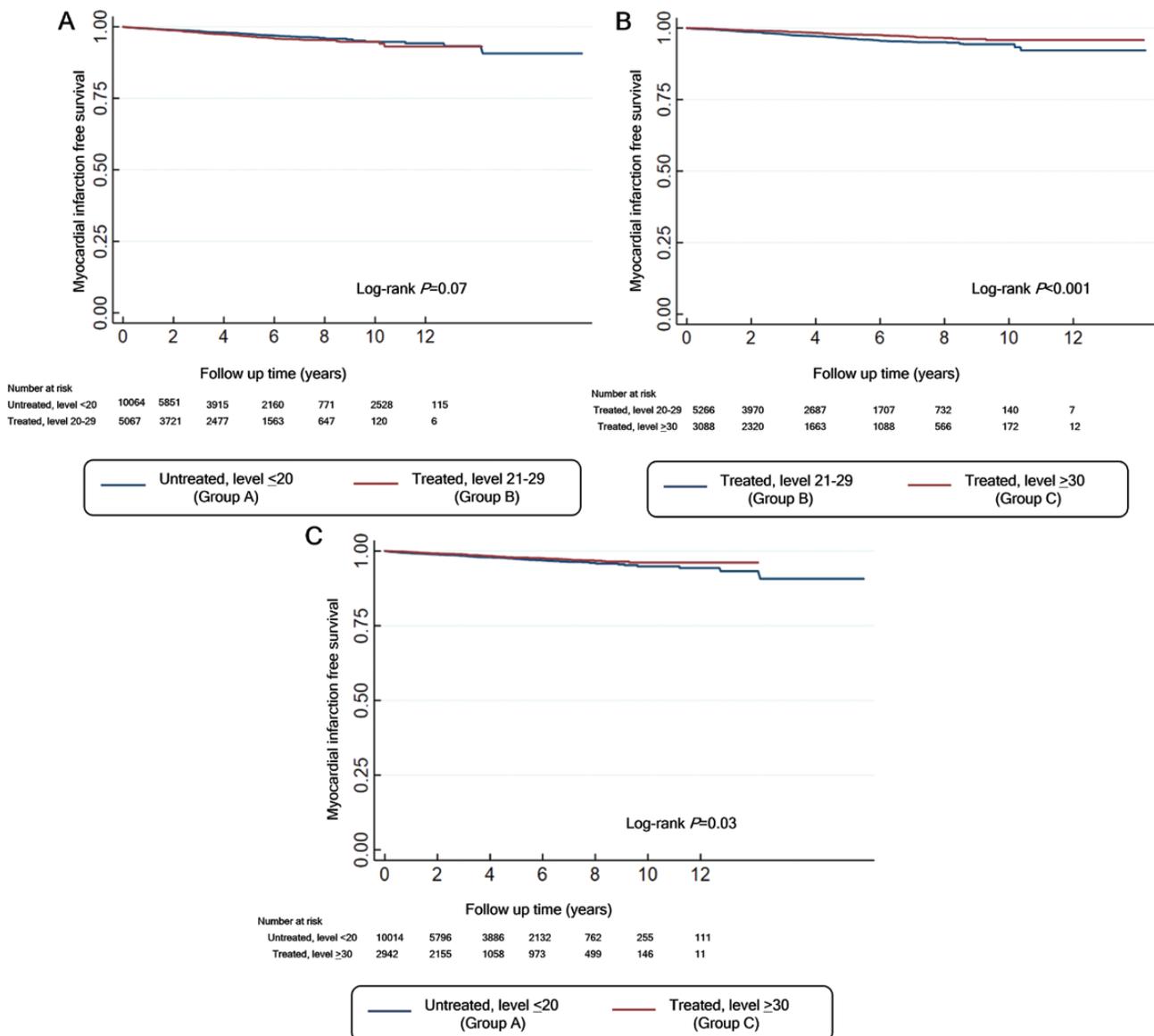


Figure 2. Kaplan–Meier curve depicting myocardial infarction (MI)–free survival among propensity-matched study groups of patients. The comparison of MI-free survival. Kaplan–Meier curves and log-rank test were utilized. Group C (treated, level ≥ 30 ng/mL group) had higher probability of MI-free survival than Group B (treated, level 20–29 ng/mL group) (log-rank $P < .001$) and Group A (untreated, level < 20 ng/mL group) (log-rank $P = .03$) group. There was no significant difference in MI free survival comparing Group A with Group B (log-rank $P = .10$).

reduction in the MI events has been reported with longer periods of observation (> 10 years) [4–6, 10, 11] compared to the clinical trials terminating at 5–6 years.

The pathophysiological mechanism for our findings remains speculative. The predominant cause of mortality in patients with (25-OH)D levels ≤ 20 ng/mL is likely multifactorial, and possibly related to the pleiotropic effect of Vit-D on immunity, cardiovascular health, and metabolic abnormalities associated with its deficiency [24–29]. Additionally, our data suggest that in Vit-D–deficient patients, post-treatment (25-OH)D levels of 21–29 ng/dL may provide inadequate protection against MI and to derive significant MI benefit post-treatment (25-OH)D levels should be > 30 ng/dL. Experimental studies have demonstrated that

Vit-D inhibits the transformation of macrophages to foam cells, increases cholesterol efflux in macrophages, improves endothelial nitric oxide formation, promotes vascular repair, and decreases thrombogenicity as well as inflammation. All these mechanisms may play a role in providing a protective effect against the atherothrombotic process such as MI [1, 27, 30].

Limitations and Strengths

This was an observational study because of which unmeasured confounding or hidden bias might be present. We were unable to account for seasonal variability in the (25-OH)D levels or the methodology utilized to measure

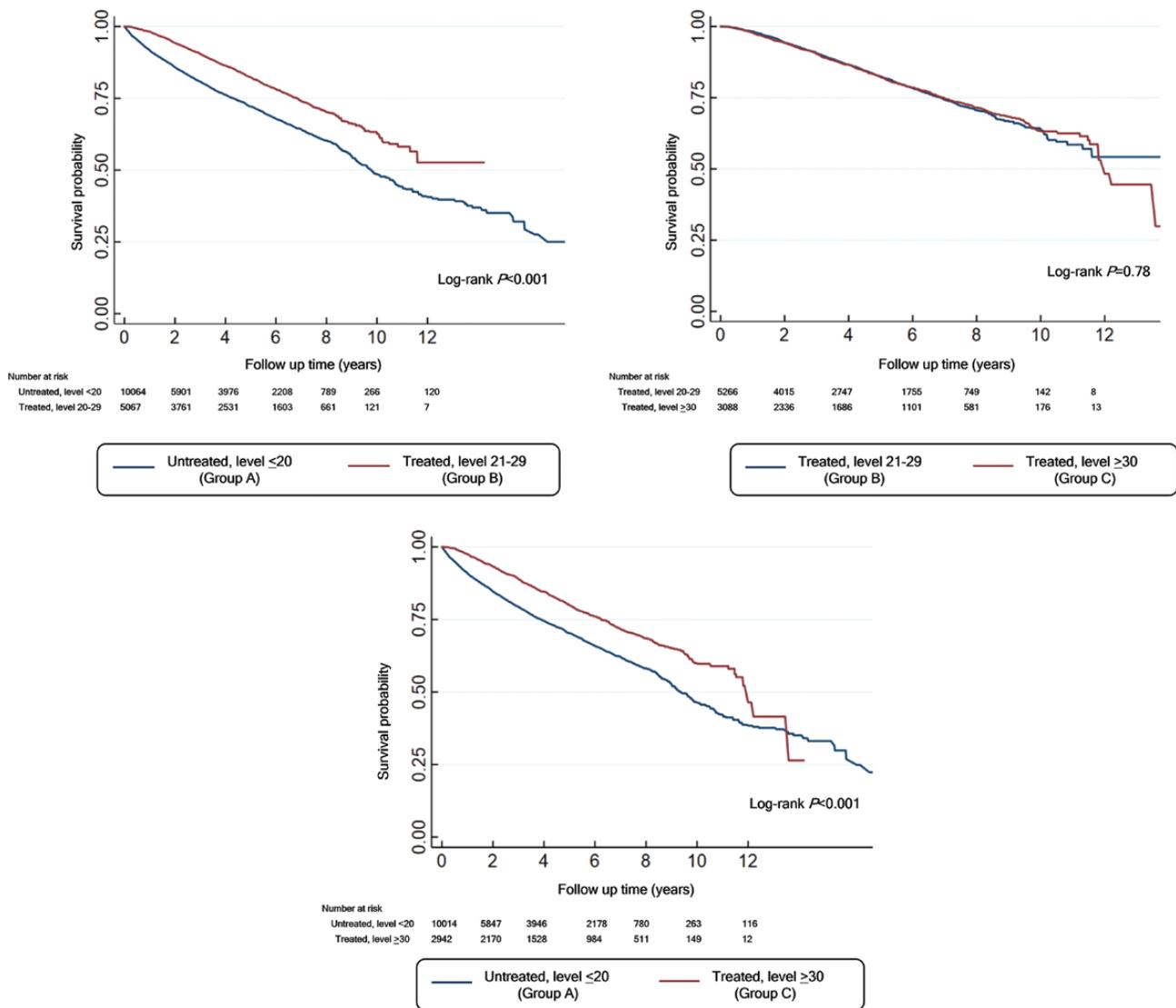


Figure 3. Kaplan–Meier curve depicting survival probability among propensity-matched study groups of patients. Comparison of survival probability between the 3 groups. Kaplan–Meier curves and log-rank test were utilized. Compared with Group A (untreated, level <20 ng/mL), the probability of survival was significantly higher in Group B (treated, level 20-29 ng/mL) (log-rank $P < .001$) and Group C (treated, level ≥ 30 ng/mL) (log-rank $P < .001$). The probability of survival was not different between Group B and Group C (log-rank $P = .78$).

(25-OH)D levels in each individual cases. We were unable to account for the use of over-the-counter Vit-D supplements that were not listed in the medical record. Our database does not have all the clinical data regarding indications for initiating and not initiating Vit-D treatment. Additionally, we were unable to determine the compliance and duration of therapy. The cause of death could not be ascertained because which cardiovascular cause of mortality was not measured. Furthermore, outcomes were determined using ICD-9 and ICD-10 codes which could have its own limitations. The results of our study may not be applicable to other populations as this study only included veterans, which is an unique population. Race variable is not available in the database available to us, hence racial differences in the population could not be accounted for.

The strength of our study is that we only included patients with low (25-OH)D levels (≤ 20 ng/mL) with extensive follow-up of up to 14 years. Each patient had at least 2 separate measurements of (25-OH)D levels to confirm the status and to measure the effect of Vit-D supplementation. We only included patients with consistent levels within each group. We were also able to stratify the patient population according to the (25-OH)D level that was maintained over the years of follow-up rather than relying only on the baseline (25-OH)D level or the dose and type of Vit-D treatment received.

Conclusion

Results from our current study suggest in patients with Vit-D deficiency and no prior history of MI, treatment to

the (25-OH)D level of >20 ng/mL was associated with a significantly lower risk of all-cause mortality. Our study also highlights that in this population reduction in the risk of MI was observed only with the increase in the (25-OH) D levels to ≥ 30 ng/mL. In the future, adequately powered, prospective, well-designed trials with a long-term follow-up will be needed to reach a conclusive agreement regarding the effect of Vit-D supplementation, and postsupplement (25-OH)D target levels on MI risk.

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Additional Information

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References

- Pilz S, Verheyen N, Gröbler MR, Tomaschitz A, März W. Vitamin D and cardiovascular disease prevention. *Nat Rev Cardiol.* 2016;**13**(7):404-417.
- Lerchbaum E, Trummer C, Theiler-Schwetz V, et al. Effects of vitamin D supplementation on body composition and metabolic risk factors in men: a randomized controlled trial. *Nutrients.* 2019;**11**(8):1894.
- Holick MF, Binkley NC, Bischoff-Ferrari HA, et al.; Endocrine Society. Evaluation, treatment, and prevention of vitamin D deficiency: an Endocrine Society clinical practice guideline. *J Clin Endocrinol Metab.* 2011;**96**(7):1911-1930.
- Barbarawi M, Kheiri B, Zayed Y, et al. Vitamin D supplementation and cardiovascular disease risks in more than 83 000 individuals in 21 randomized clinical trials: a meta-analysis. *JAMA Cardiol.* 2019;**4**(8):765-776.
- Manson JE, Cook NR, Lee IM, et al.; VITAL Research Group. Vitamin D supplements and prevention of cancer and cardiovascular disease. *N Engl J Med.* 2019;**380**(1):33-44.
- Scragg R, Stewart AW, Waayer D, et al. Effect of monthly high-dose Vitamin D supplementation on cardiovascular disease in the vitamin d assessment study: a randomized clinical trial. *JAMA Cardiol.* 2017;**2**(6):608-616.
- Bjelakovic G, Glud LL, Nikolova D, et al. Vitamin D supplementation for prevention of mortality in adults. *Cochrane Database Syst Rev.* 2014(1). CD number: CD007470.
- Brøndum-Jacobsen P, Benn M, Jensen GB, Nordestgaard BG. 25-hydroxyvitamin d levels and risk of ischemic heart disease, myocardial infarction, and early death: population-based study and meta-analyses of 18 and 17 studies. *Arterioscler Thromb Vasc Biol.* 2012;**32**(11):2794-2802.
- Autier P, Gandini S. Vitamin D supplementation and total mortality: a meta-analysis of randomized controlled trials. *Arch Intern Med.* 2007;**167**(16):1730-1737.
- Giovannucci E, Liu Y, Hollis BW, Rimm EB. 25-hydroxyvitamin D and risk of myocardial infarction in men: a prospective study. *Arch Intern Med.* 2008;**168**(11):1174-1180.
- Kestenbaum B, Katz R, de Boer I, et al. Vitamin D, parathyroid hormone, and cardiovascular events among older adults. *J Am Coll Cardiol.* 2011;**58**(14):1433-1441.
- Chowdhury R, Kunutsor S, Vitezova A, et al. Vitamin D and risk of cause specific death: systematic review and meta-analysis of observational cohort and randomised intervention studies. *BMJ.* 2014;**348**:g1903.
- Wang L, Song Y, Manson JE, et al. Circulating 25-hydroxy-vitamin D and risk of cardiovascular disease: a meta-analysis of prospective studies. *Circ Cardiovasc Qual Outcomes.* 2012;**5**(6):819-829.
- Gaksch M, Jorde R, Grimnes G, et al. Vitamin D and mortality: Individual participant data meta-analysis of standardized 25-hydroxyvitamin D in 26916 individuals from a European consortium. *PLoS One.* 2017;**12**(2):e0170791.
- Rejnmark L, Bislev LS, Cashman KD, et al. Non-skeletal health effects of vitamin D supplementation: a systematic review on findings from meta-analyses summarizing trial data. *PLoS One.* 2017;**12**(7):e0180512.
- US Department of Veterans Affairs. *VA Informatics and Computing Infrastructure (VINCI)*. 2014. https://www.hsrd.research.va.gov/for_researchers/vinci/. Accessed June 30, 2020.
- Holick MF. Vitamin D deficiency. *N Engl J Med.* 2007;**357**(3):266-281.
- Austin PC. An introduction to propensity score methods for reducing the effects of confounding in observational studies. *Multivariate Behav Res.* 2011;**46**(3):399-424.
- Austin PC, Stuart EA. Moving towards best practice when using inverse probability of treatment weighting (IPTW) using the propensity score to estimate causal treatment effects in observational studies. *Stat Med.* 2015;**34**(28):3661-3679.
- Bouillon R, Carmeliet G, Verlinden L, et al. Vitamin D and human health: lessons from vitamin D receptor null mice. *Endocr Rev.* 2008;**29**(6):726-776.
- Dobnig H, Pilz S, Scharnagl H, et al. Independent association of low serum 25-hydroxyvitamin d and 1,25-dihydroxyvitamin D levels with all-cause and cardiovascular mortality. *Arch Intern Med.* 2008;**168**(12):1340-1349.
- Daraghmeh AH, Bertoia ML, Al-Qadi MO, Abdulbaki AM, Roberts MB, Eaton CB. Evidence for the vitamin D hypothesis: the NHANES III extended mortality follow-up. *Atherosclerosis.* 2016;**255**:96-101.
- Zittermann A, Ernst JB, Prokop S, et al. Effect of vitamin D on all-cause mortality in heart failure (EVITA): a 3-year randomized clinical trial with 4000 IU vitamin D daily. *Eur Heart J.* 2017;**38**(29):2279-2286.
- Al Mheid I, Quyyumi AA. Vitamin D and cardiovascular disease: controversy unresolved. *J Am Coll Cardiol.* 2017;**70**(1):89-100.
- Tuohimaa P. Vitamin D and aging. *J Steroid Biochem Mol Biol.* 2009;**114**(1-2):78-84.
- Christakos S, Dhawan P, Verstuyf A, Verlinden L, Carmeliet G. Vitamin D: metabolism, molecular mechanism of action, and pleiotropic effects. *Physiol Rev.* 2016;**96**(1):365-408.

27. Kassi E, Adamopoulos C, Basdra EK, Papavassiliou AG. Role of vitamin D in atherosclerosis. *Circulation*. 2013;**128**(23):2517-2531.
28. Riek AE, Oh J, Darwech I, Moynihan CE, Bruchas RR, Bernal-Mizrachi C. 25(OH) vitamin D suppresses macrophage adhesion and migration by downregulation of ER stress and scavenger receptor A1 in type 2 diabetes. *J Steroid Biochem Mol Biol*. 2014;**144**(Pt A):172-179.
29. Jeng L, Yamshchikov AV, Judd SE, et al. Alterations in vitamin D status and anti-microbial peptide levels in patients in the intensive care unit with sepsis. *J Transl Med*. 2009;**7**(1):1-9.
30. Wu WX, He DR. Low vitamin D levels are associated with the development of deep venous thromboembolic events in patients with ischemic stroke. *Clin Appl Thromb Hemost*. 2018;**24**(9 Suppl):69S-75S.