Circulating Vitamin D Levels and Alzheimer's Disease: A Mendelian Randomization Study in the IGAP and UK Biobank

⁵ Longcai Wang^{a,1}, Yanchun Qiao^{b,1}, Haihua Zhang^c, Yan Zhang^b, Jiao Hua^d, Shuilin Jin^d

- ⁷ ^aDepartment of Anesthesiology, The Affiliated Hospital of Weifang Medical University, Weifang, China
- ^bDepartment of Pathology, The Affiliated Hospital of Weifang Medical University, Weifang, China
- ⁹ ^cBeijing Institute for Brain Disorders, Capital Medical University, Beijing, China
- ¹⁰ ^dDepartment of Mathematics, Harbin Institute of Technology, Harbin, China-
- ¹¹ ^eDepartment of Neurology, Xuanwu Hospital, Capital Medical University, Beijing, China
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Abstract. Observational studies strongly supported the association of low levels of circulating 25-hydroxyvitamin D (250HD) 13 and cognitive impairment or dementia in aging populations. However, randomized controlled trials have not shown clear 14 evidence that vitamin D supplementation could improve cognitive outcomes. In fact, some studies reported the association 15 between vitamin D and cognitive impairment based on individuals aged 60 years and over. However, it is still unclear that 16 whether vitamin D levels are causally associated with Alzheimer's disease (AD) risk in individuals aged 60 years and over. 17 Here, we performed a Mendelian randomization (MR) study to investigate the causal association between vitamin D levels and 18 AD using a large-scale vitamin D genome-wide association study (GWAS) dataset and two large-scale AD GWAS datasets 19 from the IGAP and UK Biobank with individuals aged 60 years and over. Our results showed that genetically increased 20 21 25OHD levels were significantly associated with reduced AD risk individuals aged 60 years and over. Hence, our findings in combination with previous literature indicate that maintaining adequate vitamin D status in older people especially aged 60 22 years and over, may contribute to slow down cognitive decline and to forestall AD. Long-term randomized controlled trials 23 are required to test whether vitamin D supplementation may prevent AD in older people especially those aged 60 years and 24 may be recommended as preventive agents. 25

26 Keywords: Alzheimer's disease, genome-wide association study, Mendelian randomization, vitamin D

Alzheimer's disease (AD) is the most common neurodegenerative disorder [1–4]. It is well known that extracellular deposition of amyloid plaques mainly consisting of amyloid- β (A β) peptide is one of the core pathological features of AD [1, 4–6]. Meanwhile, multiple lines of evidence indicate that

INTRODUCTION

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⁶ and Guiyou Liu^{c,e,*}

¹These authors contributed equally to this work.

^{*}Correspondence to: Guiyou Liu, Department of Neurology, Xuanwu Hospital, Capital Medical University, Room 1037, Donghuajinzuo, Guanganmennei Street, XiCheng District, Beijing 100053, China. E-mails: liu_gy@tib.cas.cn or liuguiyou1981@163.com

oxidative stress is involved in the pathogenesis of AD 34 [7]. Importantly, $A\beta$ is toxic in neuronal cell cul-35 tures through a mechanism involving free radicals [7]. 36 The clearance of $A\beta$ could protect against apoptosis 37 which could usually induce the oxidative stress and 38 further cause damage in the brain of AD patients [8]. 30 Evidence from animal models of AD shows that 40 vitamin D could reduce oxidative stress, prevent neu-41 rons from dying, and further mediate the clearance of 42 Aβ plaques by activating macrophages [8]. In addi-43 tion, various studies using animal models of aging and 44 AD showed that vitamin D supplementation could 45 protect against biological processes associated with 46 AD and enhances learning and memory performance 47 [9]. Meanwhile, human observational or genetic stud-48 ies have also investigated the role of vitamin D in AD, 49 and strongly supported the association of low levels of 50 circulating 25-hydroxyvitamin D (25OHD) and cog-51 nitive impairment or dementia in aging populations 52 [8-14]. 53

Importantly, Mendelian randomization studies 54 have been used to determine the causal inferences 55 and showed that genetically increased vitamin D 56 levels could reduce the risk of AD [15-17]. How-57 ever, randomized controlled trials have not shown 58 clear evidence that vitamin D supplementation could 59 improve cognitive outcomes [9, 18]. In fact, there is 60 still a controversial link between vitamin D levels and 61 cognitive performance [9]. In a recent review, Lan-62 del et al. discussed the specificity by which vitamin 63 D could improve cognitive performance in humans 64 [9]. Landel et al. suggested a possible age threshold 65 [9]. In brief, some studies reported the association 66 between vitamin D and cognitive impairment based 67 on individuals aged 60 years and over [9]. 68

Until now, it is still unclear that whether vitamin 69 D levels are causally associated with AD risk in indi-70 viduals aged 60 years and over. Here, we performed a 71 Mendelian randomization (MR) study to investigate 72 the causal association between vitamin D levels and 73 AD using a large-scale vitamin D genome-wide asso-74 ciation study (GWAS) dataset and two large-scale AD 75 GWAS datasets from the International Genomics of 76 Alzheimer's Project (IGAP) and UK Biobank with 77 individuals aged 60 years and over [19, 20]. 78

79 MATERIALS AND METHODS

80 Study design

MR is based on three principal assumptions. Here, we described these three principal assumptions using the association between vitamin D levels and AD as an example. First, the instrumental variables (genetic variants) should be significantly associated with the exposure (vitamin D levels), such as the genome-wide significant level (p < 5.00E-08) [21, 22]. Second, instrumental variables should not be associated with confounders [21, 22]. Third, instrumental variables should affect the risk of the outcome (AD) only through the exposure (vitamin D levels) [21, 22]. In general, the second and third assumptions are collectively known as independence from pleiotropy [23]. Here, MR is based on the large-scale publicly available GWAS summary datasets in vitamin D and AD. All participants have given informed consent in all these corresponding original studies [19, 20].

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Vitamin D genetic variants

We selected six genetic variants associated with circulating 25OHD levels achieving a genome-wide significant level (p < 5.00E-08) as the potential instrumental variables, which are around six loci including GC, NADSYN1/DHCR7, CYP2R1, CYP24A1, SEC23A, and AMDHD1 from a recent GWAS including 79,366 (all European descent) [24]. These six genetic variants are located at five different chromosomes (Table 1). Two genetic variants rs12785878 (chr11: 71167449) and rs10741657 (chr11:14914878) are located the same chromosome 5. However, the distance between both variants is 56252571 bp. Hence, all these six genetic variants were independent and not in linkage disequilibrium, as described in the original study [24]. Here, we provided the summary results about the effect of each genetic variant on 250HD levels and the standard errors in Table 1.

IGAP AD GWAS dataset

The AD GWAS dataset is from the large-scale 118 meta-analysis performed by the IGAP [20]. In stage 119 1, the IGAP performed a meta-analysis of 46 GWAS 120 datasets including 21,982 cases and 41,944 cogni-121 tively normal controls of European descent from four 122 consortia including the Alzheimer Disease Genetics 123 Consortium (ADGC), Cohorts for Heart and Aging 124 Research in Genomic Epidemiology Consortium 125 (CHARGE), The European Alzheimer's Disease Ini-126 tiative (EADI), and the Genetic and Environmental 127 Risk in AD/Defining Genetic, Polygenic and Envi-128 ronmental Risk for Alzheimer's Disease Consortium 129 (GERAD/PERADES) [20]. All patients with AD 130

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SNP	Position (hg19)	Nearby genes	EA/NEA	EAF	Beta	SE	р
rs3755967	chr4:72609398	GC	C/T	0.72	0.089	0.0023	4.74E-343
rs12785878	chr11:71167449	DHCR7	T/G	0.75	0.036	0.0022	3.80E-62
rs10741657	chr11:14914878	CYP2R1	A/G	0.40	0.031	0.0022	2.05E-46
rs17216707	chr20:52732362	CYP24A1	T/C	0.79	0.026	0.0027	8.14E-23
rs10745742	chr12:96358529	AMDHD1	T/C	0.40	0.017	0.0022	1.88E-14
rs8018720	chr14:39556185	SEC23A	G/C	0.18	0.017	0.0029	4.72E-09

 Table 1

 Characteristics of six genetic variants in vitamin D GWAS dataset

SNP, single-nucleotide polymorphism; EA, Effect Allele; NEA, Non-Effect Allele; EAF, Effect Allele Frequency; SE, standard error. Beta is the regression coefficient based on the vitamin D raising allele (effect allele).

satisfied the NINCDS-ADRDA criteria or DSM-IV guidelines [20, 25]. The average age at onset for all AD cases is \geq 73, and the average age at examination for 83% controls is \geq 76 [20].

135 UK Biobank AD GWAS dataset

The UK Biobank is a large national and interna-136 tional health resource, which could be used to identify 137 the causes of many complex diseases in middle aged 138 and older individuals (http://www.ukbiobank.ac.uk) 139 [26]. A total of 502,536 community-dwelling individ-140 uals aged between 37 and 73 years were recruited in 141 the United Kingdom between 2006 and 2010 [26]. 142 The proportion of women was 56% and the aver-143 age age was 56 (SD 8) in both women and men 144 [26]. Here, we selected a large GWAS of AD-by-145 proxy by analyzing 314,278 participants from the UK 146 Biobank including 27,696 maternal cases and 14,338 147 paternal cases [19]. In this GWAS dataset, a proxy 148 phenotype for AD case-control status was assessed 149 via self-report [19]. Participants were asked to report 150 "Has/did your father or mother ever suffer from 151 Alzheimer's disease/dementia?" Participants whose 152 parents were aged less than 60 years, dead before 153 reaching age 60 years, or without age information, 154 were excluded [19]. 155

156 Pleiotropy analysis

To meet MR assumptions, we performed a compre-157 hensive pleiotropy analysis to assure that the vitamin 158 D genetic variants affect AD risk not through biolog-159 ical pathways independent of vitamin D levels. For 160 the known AD risk factors, we manually evaluated the 161 association of vitamin D variants with the leading AD 162 risk factors including low levels of education, midlife 163 hearing loss, physical inactivity, high blood pressure 164 (hypertension), type 2 diabetes, obesity, smoking, 165 depression, and social isolation [27]. The significance 166 threshold for the association of these six vitamin D 167

variants with known confounders is a Bonferroni cor-168 rected significance threshold p < 0.05/6 = 0.00833. 169 Here, we provided more detailed information about 170 the manual pleiotropy analysis in the Supplementary 171 Methods. For the unknown confounders, we selected 172 two statistical methods including MR-Egger inter-173 cept test to assess the presence of potential pleiotropy 174 [28], and Mendelian randomization pleiotropy resid-175 ual sum and outlier (MR-PRESSO) test to identify 176 the horizontal pleiotropic outliers [29]. Here, we pro-177 vided more detailed information about the MR-Egger 178 intercept test method in the Supplementary Methods. 179 The threshold of statistical significance for evidence 180 of pleiotropy is p < 0.05. 181

Mendelian randomization analysis

We first adjusted the effect alleles of six vitamin D genetic variants to be associated with increased vitamin D levels in Table 1. We then transferred and further aligned the effect alleles of these six genetic variants in diagnosed AD and self-report AD-byproxy GWAS datasets to be consistent with the effect alleles of these six genetic variants in the vitamin D GWAS dataset. In these six vitamin D genetic variants, rs8018720 (G/C, G with the minor allele frequency (MAF) = 0.18) is an ambiguous palindromic variant (i.e., with alleles either A/T or C/G). Hence, we selected its proxy rs2144530 (C/T, C with the MAF = 0.18), which showed high linkage disequilibrium with rs8018720 ($r^2 = 1$ and D' = 1) using the HaploReg v4.1 based on the linkage disequilibrium information in 1000 Genomes Project (CEU) [30].

Here, suppose we have successfully extracted the summary results including beta coefficients and their standard errors about the associations of each genetic variant G_j (j=1,...,6) with vitamin D levels $(\hat{\beta}_{Xj}, se(\hat{\beta}_{Xj}))$ and AD $(\hat{\beta}_{Yj}, se(\hat{\beta}_{Yj}))$. For a given vitamin D genetic variant, the causal effect of vitamin D levels on AD can be consistently estimated

as a simple ratio of association estimate $\hat{\theta}_j = \frac{\hat{\beta}_{Y_j}}{\hat{\beta}_{X_j}}$ and its approximate variance $v_j = \frac{se(\hat{\beta}_{Y_j})^2}{\hat{\beta}_{X_j}^2}$. Here, we 206 207 selected the inverse-variance weighted meta-analysis 208 (IVW) as the main analysis to combined the variant-200 specific estimates to get the overall estimate [22]. 210 In addition, we selected other two sensitivity anal-211 vsis methods including weighted median regression 212 and MR-PRESSO [29], which could examine the 213 robustness of the estimate with each other. Here, we 214 provided more detailed information about the IVW 215 and weighted median regression methods in the Sup-216 plementary Methods. 217

Meanwhile, we conducted a leave-one-out permu-218 tation analysis by removing each genetic variant and 219 recalculating the overall effect estimate, which could 220 evaluate the influence of single genetic variant on the 221 estimate. The odds ratio (OR) as well as 95% confi-222 dence interval (CI) of AD corresponds to about each 223 genetically determined standard deviation (SD) (25 224 nmol/L) increase in natural-log transformed 25OHD 225 levels. All analyses were conducted using the R pack-226 age 'MendelianRandomization' [31]. The threshold 227 of statistical significance for the potential genetic 228 association between vitamin D levels and AD risk 229 was p < 0.05. 230

231 Power analysis

The proportion of vitamin D variance explained by 232 the six vitamin D genetic variants could be estimated 233 by \mathbb{R}^2 . It is estimated that these six vitamin D genetic 234 variants could explain about 2.84% of the 25OHD 235 variance (R^2) [24]. The strength of the six vitamin 236 D genetic variants could be evaluated using the 237 first-stage F-statistic. F > 10 could avoid bias in MR 238 studies [32]. Here, we calculated the F-statistic and 239 statistical power to estimate the minimum detectable 240 magnitudes of association using the web-based tool 241 mRnd (https://cnsgenomics.shinyapps.io/mRnd/) 242 and a two-sided type-I error rate $\alpha = 0.05$ [33]. 243

244 **RESULTS**

245 AD summary statistics

Using the six vitamin D genetic variants, we
extracted their corresponding AD and AD-by-proxy
summary statistics in the IGAP and UK Biobank
GWAS datasets, respectively, as provided in Supplementary Table 1 and Supplementary Table 2.

Pleiotropy analysis

The manual pleiotropy analysis showed that none of these six vitamin D genetic variants was significantly associated with known confounders at the Bonferroni corrected significance threshold (p < 0.05/6 = 0.00833). More detailed results are provided in Supplementary Table 3. MR-Egger intercept test showed no significant pleiotropy in the IGAP GWAS dataset (intercept = -0.001, and p = 0.927) and UK Biobank GWAS dataset (intercept = -0.012, and p = 0.086). In addition, MR-PRESSO test identified no horizontal pleiotropic outliers.

Mendelian randomization analysis

In the IGAP GWAS dataset, IVW showed that the genetically increased 25OHD levels (per 1 SD increase) were significantly associated with the reduced AD (OR=0.62, 95% CI: 0.46–0.84, p=0.002). Interestingly, two sensitivity analysis methods support the significant association of genetically increased 25OHD levels with the reduced AD with p < 0.05. The estimates from both sensitivity analysis methods were consistent with the IVW estimate in terms of direction and magnitude including weighted median (OR=0.64, 95% CI: 0.46–0.89, p=0.007) and MR-PRESSO (OR=0.62, 95% CI: 0.51–0.75, p=0.0047). Figure 1 shows individual genetic estimates from each of the 6 genetic variants in the IGAP GWAS dataset.

In the UK Biobank GWAS dataset, all these three methods showed suggestive effect of 25OHD levels on AD risk. The estimates are similar with those from the IGAP in terms of direction. However, the 95% CI included the null including IVW (OR = 0.88, 95% CI: 0.73–1.06, p=0.19), weighted median (OR = 0.94, 95% CI: 0.76–1.14, p=0.51), and MR-PRESSO (OR = 0.88, 95% CI: 0.74–1.06, p=0.25). Figure 2 shows individual genetic estimates from each of the 6 genetic variants in the UK Biobank GWAS dataset.

In both the IGAP and UK Biobank GWAS datasets, the leave-one-out permutation analysis further showed that the direction and precision of the estimates between 25OHD levels and AD remained largely unchanged using all these three methods (Table 2). All these findings suggest that our results are robust.

Power analysis

In the IGAP GWAS dataset, the first-stage Fstatistic was 1869.57 > 10. Our study had 80% power 251

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Study	TE	seTE	Odds	Ratio		OR	95%-CI	Weight (fixed)	Weight (random)
rs3755967 rs10741657 rs12785878 rs10745742	-0.47 -0.36 -0.31 -0.02	0.1809 0.4742 0.4556 0.8588			(((0.63 0.70 0.73 0.98	[0.44; 0.89] [0.28; 1.76] [0.30; 1.79] [0.18; 5.26]	69.6% 10.1% 11.0% 3.1%	69.6% 10.1% 11.0% 3.1%
rs8018720 rs17216707 Fixed effect model Random effects model	-0.81 -1.37	1.1118 0.7192		•	().44).25).62	[0.05; 3.92] [0.06; 1.04] [0.46; 0.84]	1.8% 4.4% 100.0%	1.8% 4.4%
Heterogeneity: $I^2 = 0\%$, τ^2	= 0, p =	0.83	0.1 0.5	<mark> </mark> 1 2	` 10			O	

Fig. 1. Individual genetic estimates from each of the 6 genetic variants using the IGAP GWAS dataset.

Study	ТЕ	seTE		Odds	Ratio	OR	95%-CI	Weight (fixed)	Weight (random)
rs3755967	-0.06	0.1124			ŀ	0.94	[0.75; 1.17]	70.4%	70.4%
rs10741657	0.05	0.3226				1.05	[0.56; 1.97]	8.5%	8.5%
rs12785878	-0.09	0.3000				0.92	[0.51; 1.65]	9.9%	9.9%
rs8018720	-0.80	0.5941		+	-	0.42	[0.13, 1.33]	2.5%	2.5%
rs17216707	-0.43	0.3808	-	* !	_	0.65	[0.31; 1.37]	6.1%	6.1%
Fixed effect model				4		0.88	[0 73. 1 06]	100.0%	
Random effects model				\diamond		0.88	[0.73; 1.06]		100.0%
Heterogeneity: $I^2 = 0\%$, τ^2	= 0, p =	0.42					• • •		
			0.2	0.5 1	2	5			

Fig. 2. Individual genetic estimates from each of the 6 genetic variants using the UK Biobank GWAS dataset.

Т	able 2
Leave-one-out permutation analysis of the association betwee	en 250HD levels and AD in the IGAP and UK Biobank datasets

Dataset	Excluded	IVW				Weighted me	dian	MR-PRESSO			
	SNP	OR	95% CI	p	OR	95% CI	р	OR	95% CI	р	
IGAP	rs3755967	0.62	0.36-1.06	0.08	0.71	0.37-1.35	0.30	0.62	0.42-0.91	0.07	
IGAP	rs12785878	0.62	0.45 - 0.84	2.00E-03	0.64	0.45-0.89	8.00E-03	0.62	0.49-0.77	1.29E-02	
IGAP	rs10741657	0.61	0.45 - 0.84	2.00E-03	0.63	0.45 - 0.88	7.00E-03	0.61	0.49-0.76	1.16E-02	
IGAP	rs17216707	0.61	0.45-0.83	1.00E-03	0.64	0.46-0.88	7.00E-03	0.61	0.50-0.75	9.17E-03	
IGAP	rs10745742	0.63	0.46 - 0.84	2.00E-03	0.64	0.46-0.89	8.00E-03	0.63	0.51 - 0.77	1.24E-02	
IGAP	rs8018720	0.65	0.48-0.88	5.00E-03	0.64	0.46-0.89	9.00E-03	0.65	0.58 - 0.72	1.27E-03	
UK Biobank	rs3755967	0.77	0.55-1.08	0.13	0.87	0.57-1.32	0.50	0.77	0.55 - 1.08	0.20	
UK Biobank	rs12785878	0.87	0.71-1.07	0.19	0.93	0.76-1.15	0.51	0.87	0.71 - 1.07	0.26	
UK Biobank	rs10741657	0.88	0.71-1.09	0.25	0.93	0.75-1.14	0.47	0.88	0.71-1.09	0.31	
UK Biobank	rs17216707	0.90	0.75-1.09	0.28	0.94	0.76-1.15	0.52	0.90	0.76 - 1.07	0.30	
UK Biobank	rs10745742	0.90	0.75-1.09	0.29	0.94	0.76-1.15	0.52	0.90	0.78 - 1.06	0.27	
UK Biobank	rs8018720	0.90	0.74-1.10	0.31	0.94	0.76-1.15	0.53	0.90	0.74 - 1.10	0.36	

IGAP, International Genomics of Alzheimer's Project; IVW, inverse-variance weighted meta-analysis; MR-PRESSO, Mendelian randomization pleiotropy residual sum and outlier; OR, odds ratio; CI, confidence interval.

to detect an OR of 0.87 or lower per SD (25
nmol/L) increase in circulating 25OHD levels for AD,
which are comparable with effect size that have been
observed in observational studies relating circulating
25OHD levels to risk of AD with OR = 0.80 [11],

and OR = 0.69 [12]. The N required for 80% power is 6135. Interestingly, the power is 89% to detect the genetic association between increased 25OHD levels and reduced AD risk with OR = 0.62. In the UK Biobank GWAS dataset, the first-stage F-statistic was

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9187.39 > 10. Our study had 80% power to detect OR
of 0.91 or lower per SD (25 nmol/L) increase in circulating 25OHD levels for AD. The N required for 80%
power is 160,101. Interestingly, the power is 95%
to detect the genetic association between increased
25OHD levels and reduced AD risk with OR = 0.88.

315 DISCUSSION

Until now, there has been an increased research 316 interest for observational studies, genetic association 317 studies, and randomized controlled trials exploring 318 the impact of vitamin D intake (diet and supple-319 ments) on AD, due to its roles beyond bone health 320 and calcium homeostasis [8]. Observational studies 321 have reported that vitamin D deficiency is associated 322 with an increased risk of AD [8-12]. However, ran-323 domized controlled trials have not provided strong 324 evidence that vitamin D supplementation could 325 improve cognitive outcomes [9]. Evidence shows that 326 vitamin D supplementation may improve cognitive 327 outcomes in individuals aged 60 years and over [9]. 328 Until now, it remains unclear whether there is a causal 329 association between increased vitamin D levels and 330 reduced AD risk in individuals aged 60 years and 331 over. Hence, we performed a MR study. Our main 332 analysis using IVW method showed that genetically 333 increased 25OHD levels were significantly associ-334 ated with reduced AD risk individuals aged 60 years 335 and over. Importantly, the estimates from other two 336 sensitivity analysis methods were consistent with the 337 IVW estimate in terms of direction and magnitude. 338 A leave-one-out permutation further suggested that 339 these estimates were robust. 340

341 *Comparison with randomized controlled trials*

In 2004, Dhesi et al. selected 139 ambulatory sub-342 jects with vitamin D insufficiency (aged 65 years 343 and over), and found that 25OHD levels in the treat-344 ment group increased significantly after 6 months 345 post-intervention [34]. Importantly, vitamin D sup-346 plementation could improve functional performance, 347 reaction time and balance [34]. In 2011, Stein et al. 348 first performed a pilot study of 13 AD individuals 349 aged > 60 with median Folstein Mini-Mental State 350 Examination (MMSE) score 21.5 [35]. These 13 AD 351 cases were treated with open label 3000 IU vita-352 min D2 tablets for 8 weeks, with dose adjustments 353 to maintain 25OHD 135-160 nM [35]. Their results 354 showed that the median 25OHD levels increased from 355 66 to 140 nM [35]. Median baseline AD assess-356

ment scale-cognitive subscale (ADAS-cog) was 25 357 and median improvement in ADAS-cog score was 358 6.0 points [35]. The Disability Assessment in Demen-359 tia (DAD) score increased in 11 out of 13, which 360 indicated less disability [35]. In 2011, Dean et al. 361 conducted a randomized controlled trial to investi-362 gate the effects of vitamin D supplementation on 363 cognitive and emotional functioning in 128 young 364 adults with the mean age of 21.8 years including 365 63 individuals in treatment group and 65 individ-366 uals in placebo group for 6 weeks [36]. Their 367 results showed no significant changes in working 368 memory, response inhibition, cognitive flexibility, 369 hallucination-proneness, psychotic-like experiences, 370 and ratings of depression, anxiety, or anger [36]. 371 In brief, randomized controlled trials have pro-372 vided evidence that vitamin D supplementation could 373 improve cognition in individuals aged 60 years and 374 over, but not in young adults. Hence, our find-375 ings are consistent with those from randomized 376 controlled trials. 377

Comparison with Mendelian randomization studies

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Until now, MR studies have been conducted to test 380 whether genetically vitamin D levels are associated 381 with AD [15–17]. There are three main differences 382 between our current study and previous studies. First, 383 previous MR studies selected four genetic variants 384 including rs2282679, rs12785878, rs10741657, and 385 rs6013897 [15, 16]. The effect sizes about these 386 four genetic variants on 250HD levels were esti-387 mated in the Canadian Multicentre Osteoporosis 388 Study (N=2,347) [15, 16]. Hence, compared with 389 2,347 samples, the effect of each variant on 25OHD 390 levels from 79,366 individuals will be more accurate, 391 as we used in the current study [24]. Second, previ-392 ous MR studies only selected the AD GWAS dataset 393 from the IGAP including 17,008 cases and 37,154 394 controls without any replication dataset [15-17]. 395 Here, we selected the IGAP GWAS dataset including 396 21,982 cases and 41,944 controls as the discovery 397 dataset, and the UK Biobank dataset as the repli-398 cation dataset. Third, previous studies reported that 399 genetically vitamin D levels were associated with 400 reduced risk of AD, but did not highlight the possible 401 age threshold. Here, we confirmed the age thresh-402 old that genetically vitamin D levels could reduce 403 the risk of AD in individuals aged 60 years and 404 over [9]. 405

Strengths and limitations 406

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we selected a large-scale vitamin D GWAS dataset 408 (N = 79,366), and two large-scale AD GWAS datasets 409 from the IGAP (N=63,926) and UK Biobank 410 (N=314,278). Second, both the vitamin D and AD 411 GWAS datasets include subjects of European descent, 412 which may reduce the influence on the potential asso-413 ciation caused by the population stratification. Third, 414 we selected six independent genetic variants as the 415 instruments, which may reduce the influence of link-416 age disequilibrium. Fourth, we selected multiple MR 417 methods, which may examine the robustness of the 418 estimate with each other. Fifth, we performed both 419 manual and statistical pleiotropy analyses, which may 420 reduce the risk of pleiotropy. 421

This MR study may have several strengths. First,

Meanwhile, this MR study may also have some 422 limitations. First, we could not completely rule out 423 additional confounders. Until now, it is almost impos-424 sible to fully rule out pleiotropy present in any MR 425 study [16, 23, 37]. Second, the causal association 426 between vitamin D level and AD risk may differ 427 across different ancestries. Hence, it should be further 428 evaluated in other ancestries. Third, leave-one-out 429 permutation analysis showed that none of these six 430 genetic variants could largely change the direction 431 and precision of the estimates between 250HD levels 432 and AD (Table 2). However, GC rs3755967 variant 433 could affect the significance, which indicates that 434 vitamin D-binding protein (DBP) (encoded by GC) 435 may have distinct effects on AD risk [15]. Hence, 436 future studies are required to evaluate the effect of 437 DBP on AD risk. Fourth, we observed significant 438 association in the IGAP, but not in the UK Biobank, 439 which indicates the difference between clinically 440 diagnosed AD and self-report AD-by-proxy [19]. 441

Conclusions 442

Until now, many clinical trials of therapies for 443 AD have failed, especially the double-blind, placebo-111 controlled, phase III trial involving patients with mild 445 dementia due to AD [38, 39]. Meanwhile, grow-446 ing evidence shows that vitamin D is involved the 447 development of AD and cognitive decline. Here, we 448 demonstrate that there is a direct causal association 449 between genetically increased vitamin D levels and 450 AD risk in people of European descent aged 60 years 451 and over. Hence, our findings in combination with 452 previous literatures indicate that maintaining ade-453 quate vitamin D status in older people especially 454

aged 60 years and over, may contribute to slow down cognitive decline and to forestall AD. Long-term randomized controlled trials are required to test whether vitamin D supplementation may prevent AD in older people especially those aged 60 years and may be recommended as preventive agents.

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