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# Serum 25-Hydroxyvitamin D, Parathyroid Hormone, and Mortality in Older Men 

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Context: Low 25-hydroxyvitamin D [25(OH)D] and high PTH may contribute to increased mortality risk in older adults.

Objective: The aim of the study was to test the association between 25(OH)D, PTH, and mortality in older men.

Design and Setting: The prospective Osteoporotic Fractures in Men (MrOS) study was conducted at six U.S. clinical centers.

Participants: We studied community-dwelling men at least 65 yr old ( $n=1490$ ).
Main Outcome Measure: Multivariate-adjusted proportional hazards models estimated the hazard ratio (HR) for mortality; cause of death was classified as cancer, cardiovascular, and other by central review of death certificates.

Results: During 7.3 yr of follow-up, 330 (22.2\%) participants died: 97 from cancer, 110 from cardiovascular disease, and 106 from other causes. The adjusted HR per sD decrease in 25(OH)D for all-cause mortality was 1.01 ( $95 \% \mathrm{Cl}, 0.89,1.14$ ); no association between 25(OH)D and cardiovascular or other-cause mortality was seen. Unexpectedly, lower 25(OH)D levels were modestly associated with a decreased risk of cancer mortality (adjusted HR per sD decrease, $0.80 ; 95 \% \mathrm{CI}, 0.64$, 0.99). Analyzing 25(OH)D as a categorical variable did not alter these results. Higher PTH levels (log-transformed) were associated with an increased risk of all-cause mortality (adjusted HR per SD increase, $1.15 ; 95 \% \mathrm{Cl}, 1.03,1.29$ ) and cardiovascular mortality (adjusted HR per so increase in PTH, 1.21; 95\% CI, 1.00, 1.45).

Conclusions: In contrast to previous studies, lower 25(OH)D levels were not associated with an increased risk of all-cause or cause-specific mortality in older men. Higher PTH levels were associated with a modest increase in mortality risk. (J Clin Endocrinol Metab 95: 4625-4634, 2010)

Vitamin D deficiency, defined as serum 25-hydroxyvitamin D [25 (OH)D] levels of less than $20 \mathrm{ng} / \mathrm{ml}(50$ nmol/liter), is common in elderly U.S. and European adults. Estimates of the prevalence of deficiency range widely from $20-100 \%$ of the older population (1-5). Observational studies (6-11) and a meta-analysis of randomized trials (12) have suggested that individuals with low $25(\mathrm{OH}) \mathrm{D}$ levels have a higher risk of mortality, even after adjustment for potentially confounding factors such as comorbid medical conditions, although some reports suggest that no association exists (13-15). Few vitamin D outcome studies have considered PTH $(14,16)$.

Individuals with excessively high levels of PTH that accompany primary or secondary hyperparathyroidism, including patients with end stage renal disease, have a high risk of cardiovascular disease and mortality (17, 18). There are few reports about PTH levels and risk of mortality in older adults without kidney disease. One reported an association between higher PTH levels and increased risk of mortality among institutionalized older adults who did not have primary hyperparathyroidism (16); another demonstrated an increased risk of all-cause and cardiovascular mortality with higher PTH levels among community dwelling men (19). The relation between PTH levels and other causes of mortality, such as cancer, have not been extensively studied.

The aim of the present analyses was to test the hypothesis that lower levels of serum $25(\mathrm{OH}) \mathrm{D}$ and higher levels of PTH were independently associated with a higher risk of all-cause and cause-specific mortality in older, commu-nity-dwelling men participating in the Osteoporotic Fractures in Men (MrOS) study.

## Subjects and Methods

The MrOS study was designed to understand healthy aging in older men with a particular focus on osteoporosis (20). Briefly, 5995 community-dwelling men aged 65 yr or older were recruited from six U.S. clinical centers (Birmingham, AL; Minneapolis, MN; Palo Alto, CA; the Monongahela Valley near Pittsburgh, PA; Portland, OR; and San Diego, CA) for a baseline visit between March 2000 and April 2002. Men must have been able to walk without assistance, not have had bilateral hip replacements, and have provided written informed consent. The study was approved by institutional review boards at all institutions.

## Assessment of serum 25(OH)D

At the baseline examination, fasting serum was collected and stored at -70 C . Season of blood draw was recorded. Serum assays of $25(\mathrm{OH}) \mathrm{D} 2$ (derived from ergocalciferol) and $25(\mathrm{OH}) \mathrm{D} 3$ (derived from cholecalciferol) were performed at the Mayo Clinic using mass spectrometry (21). $25(\mathrm{OH}) \mathrm{D} 2$ and $25(\mathrm{OH}) \mathrm{D} 3$ were summed for total $25(\mathrm{OH}) \mathrm{D}$ level. The minimum detectable limit for $25(\mathrm{OH}) \mathrm{D} 2$ was $4 \mathrm{ng} / \mathrm{ml}$ and for
$25(\mathrm{OH}) \mathrm{D} 3$ was $2 \mathrm{ng} / \mathrm{ml}$. The interassay coefficient of variation (CV) was $4.4 \%$, and the intraassay CV was $4.9 \%$.

## PTH assessment

Total intact PTH was assessed at Columbia University using the immunoradiometric assay (catalog no. 3KG600) from Scantibodies Laboratory Inc. (Santee, CA). This assay employs a polyclonal 1-84 PTH antibody with a tendency to bind in the N-terminal region of 1-84 PTH (label antibody) and a polyclonal 1-84 PTH antibody with a tendency to bind in the C-terminal region of 1-84 PTH (capture antibody). The use of these antibodies guarantees that both whole PTH (1-84 PTH) and truncated PTH fragments are detected; the assay detects both PTH 1-84 and some C-terminal fragments. Using pooled serum, the interassay CV was $8.4 \%$, and the intraassay CV was $5.6 \%$. The normal range of this assay extends to $66 \mathrm{ng} / \mathrm{ml}$. We have also measured PTH with the Scantibodies whole PTH IRMA assay (an assay that detects only PTH 1-84). The two assays yielded results that were highly correlated (Spearman correlation coefficient $=0.88 ; P<0.001$ ), and data analyses using intact and whole assay results yielded essentially the same results; thus, only the results for total intact PTH are reported here. Serum calcium was assayed by the Oregon Veterans Administration Clinical Laboratory using a Roche COBAS Integra 800 automated analyzer (Roche Diagnostics, Indianapolis, IN) on previously thawed serum. The interassay CV was $2.6 \%$.

## Mortality

Every 4 months after the baseline exam, MrOS men were contacted to provide information about fracture status. Next-of-kin were contacted when men did not return questionnaires and could not be reached by telephone. Death certificates and discharge summaries where available and were collected for all deaths reported in MrOS. Participants were followed for vital status for an average of 7.3 yr after the baseline examination, through August 2009. Vital status information is more than $99 \%$ complete. Date and cause of death from death certificates were reviewed centrally by a physician adjudicator, with cause of death classified by ICD9 codes. Cause of death categories were broadly defined as cardiovascular (ICD9 codes 394.9, 396.9, 401.1, 401.9 to $442.0,443.9,785.51,996.71$ ), cancer (ICD9 codes 150.9 to 208.0), and other causes (other recorded ICD9 codes not in the previous categories).

## Other clinical measures

Participants completed a self-administered questionnaire and clinical examination. Participants self-reported a history of cancer, stroke, heart attack, chronic obstructive pulmonary disease, hypertension, congestive heart failure, thyroid disease, Parkinson's disease, and diabetes. Height was measured using wallmounted stadiometers, and weight was measured with balance beam and digital scales. Body mass index (BMI) was calculated as weight (in kilograms) divided by height (in meters) ${ }^{2}$. Participants self-reported marital status (married $v s$. not), smoking status (current $v s$. never $v s$. past), race (Caucasian $v s$. non-Caucasian), education (college education $v s$. not), and physical activity [using the Physical Activity Scale for the Elderly (PASE)] (22). Alcohol intake was determined in a clinic interview and was classified as none/very low ( $<12$ drinks/year), light ( $\geq 12$ drinks/ year to six drinks/week), and moderate/high ( $\geq 7$ drinks/week). A functional status limitation was defined as having any diffi-
culty with the following: heavy housework, shopping, or preparing meals. A mobility limitation was defined as having any difficulty walking two or three blocks or climbing 10 stairs. A modified version of the Block Food Frequency Questionnaire was used to determine total daily dietary (supplements + food) vitamin D and calcium intake (23). Glomerular filtration rate (GFR) was estimated by the Modification of Diet in Renal Disease (MDRD) study equation (24). Percentage body fat was determined from whole body dual-energy x-ray absorptiometry (DXA) scans on Hologic 4500 scanners (Hologic, Waltham, MA). Centralized quality control procedures, certification of DXA operators, repeat scanning of phantoms and standardized procedures for scanning were used to ensure reproducibility of DXA measurements. Variability across clinics was within acceptable limits, and cross-calibration correction factors were not required.

## Participants included in analysis

Of the men enrolled at baseline, 5908 had at least one tube of serum available for analysis. From these men, 1594 were randomly selected to have serum $25(\mathrm{OH})$ D and PTH levels assayed. Of these, three were excluded from analyses due to assay problems [one with $25(\mathrm{OH})$ D levels $>3$ sD above the mean ( 75.6 $\mathrm{ng} / \mathrm{ml}$ ); one with insufficient sample for the vitamin D assay; and one with a CV for PTH assay run $>25 \%$ ]. Another 101 participants with missing data on covariates (GFR, serum calcium, phosphate, percentage body fat, or self-rated health) were excluded, leaving 1490 men for the present analyses. Cause of death analyses excluded 17 men for whom cause of death adjudication was pending as of August 2009.

## Statistical analysis

Characteristics of the MrOS participants were compared across quartiles of total serum $25(\mathrm{OH}) \mathrm{D}$ and PTH. ANOVA was used for normally distributed continuous variables, KruskalWallis for skewed continuous variables, and $\chi^{2}$ tests for categorical variables. The Spearman correlation coefficient was calculated between PTH and total $25(\mathrm{OH})$ D levels.

Total $25(\mathrm{OH}) \mathrm{D}$ levels were analyzed as quartiles and as clinical categories as suggested by Holick [normal ( $\geq 30 \mathrm{ng} / \mathrm{ml}$ ), insufficient ( $>20$ to $<30 \mathrm{ng} / \mathrm{ml}$ ), or deficient ( $\leq 20 \mathrm{ng} / \mathrm{ml}$ ) (25)], and as a continuous variable with the hazard ratio (HR) reported per SD decrease. Total intact PTH was analyzed as quartiles. To investigate the influence of clinically high PTH levels on risk of mortality, PTH was also analyzed as a dichotomous variable ( $266 \mathrm{pg} / \mathrm{ml} v$ s. $<66 \mathrm{pg} / \mathrm{ml}$ ) (26). PTH was skewed, so PTH was log-transformed for analysis as a continuous variable; results are presented per SD increase in log-transformed PTH.

To account for seasonal and geographic variation in serum $25(\mathrm{OH}) \mathrm{D}$ levels, all models were adjusted for clinical center and season of blood draw, with seasons classified as winter (JanuaryMarch), spring (April-June), summer (July-September), and fall (October-December). Cox proportional hazard models were used to estimate the HR and $95 \%$ confidence intervals (CIs) for mortality. The assumption of proportionality was tested and valid in all models. Base models that adjusted for age, season of blood draw, and clinical center were run for all-cause and causespecific mortality for PTH and $25(\mathrm{OH})$ D levels separately. Covariates that were significantly associated with serum PTH or $25(\mathrm{OH}) \mathrm{D}$ levels at the $P<0.10$ level and were associated with mortality in age-adjusted models were then included in the mul-
tivariate models. Finally, both PTH and $25(\mathrm{OH})$ D were included in a single multivariate model. Variables measuring similar characteristics (for example, percentage body fat and BMI) were selected by picking the variable or set of variables that were most strongly associated with serum PTH and 25(OH)D levels. Adjusted survival plots were also created to graphically describe the association between PTH, vitamin D, and mortality.

To test whether the association between PTH and mortality varied by $25(\mathrm{OH}) \mathrm{D}$ level (and vice versa), the significance of several interaction terms was tested. Interaction terms were evaluated for PTH $* 25(\mathrm{OH}) \mathrm{D}$ as continuous variables; $\mathrm{PTH} * 25(\mathrm{OH}) \mathrm{D}$ as quartiles; and $\mathrm{PTH} * 25(\mathrm{OH}) \mathrm{D}$ as clinical categories. Analyses were also performed that excluded men with poor renal function (estimated $\left.\mathrm{GFR}<60 \mathrm{ml} / \mathrm{min} / 1.73 \mathrm{~m}^{2}\right)$. Finally, to determine whether the effect of $25(\mathrm{OH})$ D or PTH varied over the course of the follow-up period, we completed several sensitivity analyses. We truncated follow-up time at 3 yr and at 5 yr , and we also excluded participants who died in the year after follow-up. The results of these sensitivity analyses for length of follow-up and GFR did not differ from the main analyses, so only the main analyses are reported.

## Results

Most MrOS participants were either vitamin D deficient [25 (OH)D levels $<20 \mathrm{ng} / \mathrm{ml} ; \mathrm{n}=376(25.2 \%)$ ] or insufficient [25(OH)D levels of 20 to $<30 \mathrm{ng} / \mathrm{ml} ; \mathrm{n}=737$ (49.5\%)]. Comparisons of characteristics of the MrOS cohort by quartile of vitamin D status are reported in Table 1 . Briefly, men in the lowest quartiles of $25(\mathrm{OH}) \mathrm{D}$ were older, were more likely to be nonwhite, were more likely to report functional or mobility limitations, had greater adiposity, were more often unmarried, and were somewhat more likely to report at least one comorbid medical condition or abstention from alcohol than men with higher $25(\mathrm{OH}) \mathrm{D}$ levels.

Men in the higher quartiles of PTH were older; had greater adiposity; and were more likely to report functional or mobility limitations, being unmarried, mobility limitations, and more comorbid medical conditions than men in the lower PTH quartiles (Table 2). Men with higher PTH levels also had lower vitamin D and calcium intake, lower levels of serum calcium, and worse GFR than men with lower PTH levels ( $P<0.05$ for all). PTH levels and vitamin D status were related, and the Spearman correlation coefficient between total $25(\mathrm{OH}) \mathrm{D}$ and PTH was $-0.22(P<0.001)$.

Over an average of 7.3 yr of follow-up (SD, 1.9 yr ; range, 14 d to 9.3 yr$), 330$ ( $22.2 \%$ ) participants died. Of these, $97(29.4 \%)$ were classified as cancer deaths, 110 ( $33.3 \%$ ) as cardiovascular deaths, and 106 (32.1\%) as noncancer, noncardiovascular deaths, and 17 ( $5.2 \%$ ) were unclassified. Of the 97 men who died from cancer, the most common type of cancer death was lung cancer ( $\mathrm{n}=$ $33)$. Of the 110 men with death attributed to cardio-

TABLE 1. Characteristics of the MrOS study population by quartile of $25(\mathrm{OH}) \mathrm{D}$

| Characteristic | $\begin{gathered} \text { Quartile } 1 \\ (<20.0 \mathrm{ng} / \mathrm{ml}) \end{gathered}$ | Quartile 2 (20.0 to <25.2 ng/ml) | Quartile 3 <br> (25.2 to <30 ng/ml) | Quartile 4 ( $\geq 30 \mathrm{ng} / \mathrm{ml}$ ) | $P$ value |
| :---: | :---: | :---: | :---: | :---: | :---: |
| n | 372 | 370 | 372 | 376 |  |
| Age (yr) | $74.5 \pm 6.3$ | $73.8 \pm 5.8$ | $73.9 \pm 5.7$ | $72.7 \pm 5.4$ | 0.001 |
| Weight (kg) | $84.5 \pm 14.3$ | $84.0 \pm 13.7$ | $82.4 \pm 12.2$ | $82.0 \pm 11.1$ | 0.072 |
| BMI ( $\mathrm{kg} / \mathrm{m}^{2}$ ) | $28.1 \pm 4.2$ | $27.6 \pm 3.8$ | $27.1 \pm 3.5$ | $26.8 \pm 3.1$ | <0.001 |
| Percentage body fat | $26.9 \pm 5.5$ | $26.8 \pm 5.4$ | $26.0 \pm 4.8$ | $25.7 \pm 5.1$ | 0.001 |
| Married | 283 (76.1) | 321 (86.8) | 314 (84.4) | 318 (84.6) | 0.001 |
| Nonwhite | 62 (16.7) | 30 (8.1) | 25 (6.7) | 22 (5.9) | $<0.001$ |
| Excellent/good health status | 301 (80.9) | 311 (84.1) | 321 (86.3) | 341 (90.7) | 0.002 |
| At least one functional limitation | 68 (18.3) | 40 (10.8) | 52 (14.0) | 39 (10.4) | 0.005 |
| At least one mobility limitation | 73 (19.6) | 40 (10.8) | 50 (13.4) | 37 (9.4) | $<0.001$ |
| Activity level (PASE score) | $137.5 \pm 73.3$ | $147.0 \pm 67.8$ | $147.8 \pm 64.8$ | $156.6 \pm 67.7$ | 0.002 |
| Current smoker | 22 (5.9) | 16 (4.3) | 10 (2.7) | 8 (2.1) | 0.029 |
| Alcohol use |  |  |  |  | 0.006 |
| None/very light (<12 drinks/yr) | 135 (36.3) | 130 (35.2) | 118 (31.7) | 104 (27.7) |  |
| Light ( $\geq 12$ drinks/yr to 6 drinks/wk) | 103 (27.7) | 87 (23.6) | 96 (25.8) | 132 (35.1) |  |
| Moderate ( $\geq 7$ drinks/wk) | 134 (36.0) | 152 (41.2) | 158 (42.5) | 140 (37.2) |  |
| At least one medical condition ${ }^{\text {a }}$ | 268 (72.0) | 245 (66.2) | 274 (73.7) | 252 (67.0) | 0.068 |
| Daily vitamin D intake (IU) | $272.0 \pm 227.5$ | $386.6 \pm 239.5$ | $433.9 \pm 246.8$ | $450.6 \pm 237.5$ | $<0.002$ |
| Serum calcium (mg/dl) | $9.3 \pm 0.4$ | $9.3 \pm 0.4$ | $9.3 \pm 0.4$ | $9.3 \pm 0.4$ | 0.014 |
| Albumin corrected calcium (mg/dl) | $9.5 \pm 0.5$ | $9.5 \pm 0.5$ | $9.5 \pm 0.5$ | $9.6 \pm 0.5$ | 0.001 |
| GFR ( $\mathrm{ml} / \mathrm{min} / 1.73 \mathrm{~m}^{2}$ ) | $78.9 \pm 19.4$ | $77.5 \pm 18.4$ | $75.9 \pm 17.8$ | $73.8 \pm 17.2$ | 0.017 |
| Total intact PTH level ( $\mathrm{pg} / \mathrm{ml}$ ) | $39.5 \pm 45.8$ | $33.2 \pm 13.7$ | $31.3 \pm 11.8$ | $29.7 \pm 13.0$ | <0.001 |

Data are expressed as mean $\pm$ SD or number (percentage).
${ }^{a}$ At least one of the following medical conditions: stroke, heart attack, nonskin cancer, chronic obstructive pulmonary disease, hypertension, congestive heart failure, thyroid disease, diabetes, or Parkinson's disease.
vascular disease, the most common subtype was coronary heart disease ( $\mathrm{n}=55$ ).

## 25(OH)D and all-cause mortality

There was no significant association between $25(\mathrm{OH}) \mathrm{D}$ levels and risk of all-cause mortality in this cohort (Table 3 and Fig. 1), whether $25(\mathrm{OH}) \mathrm{D}$ was analyzed as quartiles, as clinical categories, or as a continuous variable. Adjustment for potentially confounding factors did not alter this null association.

## 25(OH)D and cause-specific mortality

There was a suggestion of an association between higher levels of $25(\mathrm{OH}) \mathrm{D}$ and increased risk of cancer mortality when $25(\mathrm{OH}) \mathrm{D}$ was analyzed as clinical categories, as quartiles, or as a continuous variable. No single covariate in the fully adjusted multivariate model changed the HR for mortality [expressed per SD decrease in $25(\mathrm{OH}) \mathrm{D}$ ] by more than $4 \%$, indicating that the change in the HR from the minimally adjusted to the fully adjusted model is not due to the effects of a single variable.

There was no association between $25(\mathrm{OH}) \mathrm{D}$ levels analyzed as quartiles or clinical categories and cardiovascular or noncardiovascular, noncancer mortality in either
base or multivariate models. However, in base-adjusted models, when $25(\mathrm{OH}) \mathrm{D}$ was analyzed as a continuous variable, there was a modest increase in risk of CVD mortality per SD decrease in $25(\mathrm{OH}) \mathrm{D}(\mathrm{HR}, 1.23 ; 95 \% \mathrm{CI}$, $1.00,1.52)$. The HR was essentially unchanged after multivariate adjustment; however, this association was no longer significant.

Inclusion or exclusion of PTH from the multivariate models did not substantially alter the $25(\mathrm{OH}) \mathrm{D}$ results.

## PTH and all-cause mortality

In the base-adjusted model (adjusted for age, clinical site, and season of blood draw), the risk of mortality increased as the quartile of PTH increased (Table 4 and Fig. 1). Further adjustment for numerous potentially confounding factors, including total $25(\mathrm{OH}) \mathrm{D}$, did not substantially attenuate the association, and the association between PTH and mortality remained significant when PTH was analyzed as a continuous variable or using clinical categories $(<66 \mathrm{pg} / \mathrm{ml} v s . \geq 66 \mathrm{pg} / \mathrm{ml})$.

## PTH and cause-specific mortality

There was a suggestion of ansociation between increased PTH levels and increased risk of cardiovascular

TABLE 2. Characteristics of the MrOS study population by quartile of total intact serum PTH

| Characteristic | $\begin{gathered} \text { Quartile } 1 \\ (<23.6 \mathrm{pg} / \mathrm{ml}) \end{gathered}$ | Quartile 2 ( 23.6 to $<29.5 \mathrm{pg} / \mathrm{ml}$ ) | Quartile 3 ( 29.5 to $<38.3 \mathrm{pg} / \mathrm{ml}$ ) | $\begin{gathered} \text { Quartile } 4 \\ (\geq 38.3 \mathrm{pg} / \mathrm{ml}) \end{gathered}$ | $P$ value |
| :---: | :---: | :---: | :---: | :---: | :---: |
| n | 372 | 373 | 372 | 372 |  |
| Age (yr) | $73.3 \pm 5.6$ | $72.8 \pm 5.3$ | $73.6 \pm 5.8$ | $75.3 \pm 6.5$ | $<0.001$ |
| Weight (kg) | $81.9 \pm 11.9$ | $83.3 \pm 13.3$ | $83.8 \pm 12.7$ | $83.8 \pm 13.8$ | 0.228 |
| $\mathrm{BMI}\left(\mathrm{kg} / \mathrm{m}^{2}\right)$ | $26.9 \pm 3.3$ | $27.3 \pm 3.9$ | $27.6 \pm 3.7$ | $27.7 \pm 3.9$ | 0.006 |
| Percentage body fat | $25.86 \pm 5.1$ | $26.08 \pm 5.5$ | $26.64 \pm 5.1$ | $26.82 \pm 5.1$ | 0.038 |
| Married | 324 (87.1) | 314 (84.2) | 297 (79.8) | 301 (80.7) | 0.032 |
| Nonwhite | 30 (8.1) | 28 (7.5) | 39 (10.5) | 42 (11.3) | 0.222 |
| Excellent/good health status | 320 (86.0) | 319 (85.5) | 318 (85.5) | 317 (85.0) | 0.984 |
| At least one functional limitation | 36 (9.7) | 39 (10.5) | 54 (14.5) | 70 (18.8) | $<0.001$ |
| At least one mobility limitation | 40 (10.8) | 38 (10.2) | 61 (16.4) | 61 (16.4) | 0.011 |
| Activity level (PASE score) | $151.0 \pm 69.2$ | $151.6 \pm 68.2$ | $147.6 \pm 69.2$ | $138.9 \pm 67.8$ | 0.017 |
| Current smoker | 15 (4.0) | 15 (4.0) | 17 (4.6) | 9 (2.4) | 0.442 |
| Alcohol use |  |  |  |  |  |
| None/very light ( $<12$ drinks/yr) | 117 (31.5) | 109 (29.2) | 119 (32.0) | 142 (38.2) | 0.071 |
| Light ( $\geq 12$ drinks/yr to 6 drinks/wk) | 117 (31.5) | 107 (28.7) | 109 (29.3) | 85 (22.9) |  |
| Moderate ( $\geq 7$ drinks/wk) | 138 (37.1) | 157 (42.1) | 144 (38.7) | 145 (39.0) |  |
| At least one medical condition ${ }^{\text {a }}$ | 255 (68.6) | 248 (66.5) | 261 (70.2) | 275 (73.7) | 0.175 |
| Daily vitamin D intake (IU) | $425.5 \pm 241.3$ | $396.5 \pm 253.5$ | $385.6 \pm 252.2$ | $335.7 \pm 235.8$ | <0.001 |
| Daily calcium intake (mg) | $1261.2 \pm 617.0$ | $1166.7 \pm 632.3$ | $1101.1 \pm 565.6$ | $1047.6 \pm 550.8$ | $<0.001$ |
| Serum calcium (mg/dl) | $9.4 \pm 0.3$ | $9.3 \pm 0.4$ | $9.3 \pm 0.4$ | $9.3 \pm 0.5$ | $<0.001$ |
| Albumin corrected calcium ( $\mathrm{mg} / \mathrm{dl}$ ) | $9.6 \pm 0.4$ | $9.6 \pm 0.4$ | $9.5 \pm 0.5$ | $9.5 \pm 0.6$ | $<0.001$ |
| GFR ( $\mathrm{ml} / \mathrm{min} / 1.73 \mathrm{~m}^{2}$ ) | $79.1 \pm 16.3$ | $79.8 \pm 16.9$ | $77.1 \pm 17.9$ | $70.0 \pm 20.2$ | $<0.001$ |
| Total 25(OH)D level (ng/ml) | $27.6 \pm 8.07$ | $25.7 \pm 7.36$ | $24.7 \pm 7.55$ | $22.9 \pm 8.22$ | <0.001 |
| Vitamin D status |  |  |  |  |  |
| Deficient ( $\leq 20 \mathrm{ng} / \mathrm{ml}$ ) | 18 (4.8) | 26 (7.0) | 36 (9.7) | 67 (18.0) | $<0.001$ |
| Insufficient ( $>20$ to $<30$ $\mathrm{ng} / \mathrm{ml}$ ) | 221 (59.4) | 248 (66.5) | 264 (71.0) | 238 (63.8) |  |
| Sufficient ( $\geq 30 \mathrm{ng} / \mathrm{ml}$ ) | 133 (35.8) | 99 (26.5) | 72 (19.4) | 68 (18.2) |  |

Data are expressed as mean $\pm$ SD or number (percentage).
${ }^{a}$ At least one of the following medical conditions: stroke, heart attack, nonskin cancer, chronic obstructive pulmonary disease, hypertension, congestive heart failure, thyroid disease, diabetes, or Parkinson's disease.
mortality. In base-adjusted models, men in the highest quartile of PTH had a 2 -fold increase in the risk of cardiovascular mortality compared with those in the lowest PTH quartile, but this association was no longer significant after adjustment for potentially confounding factors. In both base-adjusted and multivariate-adjusted models, each sD increase in log-transformed PTH was associated with a modestly increased risk of cardiovascular mortality. Men with clinically elevated PTH levels ( $\geq 66 \mathrm{ng} / \mathrm{ml}$ ) had an elevated risk of cardiovascular mortality in base-adjusted models, but this association was attenuated and no longer significant in multivariate models.

In base-adjusted and multivariate-adjusted models, there was no general association with cancer mortality or noncancer, noncardiovascular mortality when PTH was analyzed as quartiles, as a dichotomous variable ( $<66 \mathrm{pg} / \mathrm{nl} v \mathrm{~s} . \geq 66 \mathrm{pg} /$ nl ), or as a log-transformed continuous variable. The only exception was when PTH was analyzed as a continuous vari-
able in multivariate-adjusted models because the HR per SD increase in log-transformed PTH reached statistical significance (HR, 1.15; 95\% CI, 1.03, 1.29).

When the analyses were restricted to participants with PTH values below $66 \mathrm{ng} / \mathrm{ml}(\mathrm{n}=1439)$, the association between PTH and all-cause mortality was attenuated and no longer significant (HR, 1.10; $95 \% \mathrm{CI}, 0.97,1.25$ ), suggesting that individuals with PTH values outside the normal range are at the highest risk of death.

## Interaction between PTH and 25(OH)D

There was no evidence for a statistical interaction between $25(\mathrm{OH}) \mathrm{D}$ and PTH levels and risk of all-cause or cause-specific mortality. Interaction terms were tested for continuous variables of PTH and $25(\mathrm{OH}) \mathrm{D}$, quartiles of PTH and $25(\mathrm{OH}) \mathrm{D}$, and clinical categories of PTH and $25(\mathrm{OH}) \mathrm{D}$. None reached the conservative $P<0.1$ level of significance.
TABLE 3. HRs $(95 \% \mathrm{Cls})$ for mortality by categories of total $25(\mathrm{OH}) \mathrm{D}$ in the MrOS study

|  | All-cause mortality |  | Cancer |  | Cardiovascular |  | Noncancer, noncardiovascular |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Base model ${ }^{\text {a }}$ | Multivariate ${ }^{\text {b }}$ | Base model ${ }^{\text {a }}$ | Multivariate ${ }^{\text {b }}$ | Base model ${ }^{\text {a }}$ | Multivariate ${ }^{\text {b }}$ | Base model ${ }^{\text {a }}$ | Multivariate ${ }^{\text {b }}$ |
| Quartiles |  |  |  |  |  |  |  |  |
| $\begin{aligned} & \text { Q1, }<19.9 \mathrm{ng} / \mathrm{ml} \\ & (\mathrm{n}=372) \end{aligned}$ | 1.11 (0.81, 1.52) | 0.95 (0.68, 1.34) | 0.63 (0.35, 1.14) | 0.52 (0.27, 1.00) | 1.50 (0.86, 2.63) | 1.52 (0.83, 2.80) | 1.26 (0.73, 2.17) | 0.94 (0.51, 1.72) |
| $\begin{aligned} & \mathrm{Q} 2, \geq 19.9 \mathrm{to} \\ & <25.2 \mathrm{ng} / \mathrm{ml} \\ & (\mathrm{n}=370) \end{aligned}$ | 1.01 (0.73, 1.39) | 1.05 (0.75, 1.47) | 0.92 (0.54, 1.58) | 0.90 (0.51, 1.60) | 1.08 (0.59, 1.97) | 1.21 (0.65, 2.28) | 1.02 (0.58, 1.81) | 1.03 (0.56, 1.87) |
| $\begin{aligned} & \mathrm{Q} 3, \geq 25.2 \mathrm{to} \\ & <30.0 \mathrm{ng} / \mathrm{ml} \\ & (\mathrm{n}=372) \end{aligned}$ | 0.94 (0.68, 1.30) | 0.89 (0.64, 1.24) | 0.82 (0.47, 1.43) | 0.80 (0.45, 1.41) | 1.18 (0.65, 2.12) | 1.12 (0.61, 2.06) | 0.88 (0.49, 1.59) | 0.80 (0.44, 1.47) |
| $\begin{gathered} \mathrm{Q} 4, \geq 30 \mathrm{ng} / \mathrm{ml} \\ (\mathrm{n}=376) \end{gathered}$ | 1.00 (referent) | 1.00 (referent) | 1.00 (referent) | 1.00 (referent) | 1.00 (referent) | 1.00 (referent) | 1.00 (referent) | 1.00 (referent) |
| $P$ for trend | 0.405 | 0.961 | 0.192 | 0.086 | 0.177 | 0.153 | 0.302 | 0.906 |
| Clinical categories |  |  |  |  |  |  |  |  |
| Deficient, $<20$ ng/ml ( $\mathrm{n}=376$ ) | 1.10 (0.81, 1.51) | 0.94 (0.67, 1.32) | 0.62 (0.34, 1.13) | 0.51 (0.27, 0.98) | 1.49 (0.85, 2.62) | 1.51 (0.82, 2.76) | 1.25 (0.72, 2.16) | 0.93 (0.51, 1.70) |
| Insufficient, 20 to $<30 \mathrm{ng} / \mathrm{ml}$ ( $\mathrm{n}=737$ ) | 0.98 (0.74, 1.29) | 0.97 (0.72, 1.30) | 0.88 (0.55, 1.41) | 0.85 (0.52, 1.40) | 1.13 (0.67, 1.92) | 1.17 (0.67, 2.02) | 0.95 (0.57, 1.58) | 0.91 (0.53, 1.53) |
| Sufficient, $\geq 30$ $\mathrm{ng} / \mathrm{ml}(\mathrm{n}=377)$ | 1.00 (referent) | 1.00 (referent) | 1.00 (referent) | 1.00 (referent) | 1.00 (referent) | 1.00 (referent) | 1.00 (referent) | 1.00 (referent) |
| $P$ for trend | 0.494 | 0.706 | 0.120 | 0.044 | 0.132 | 0.167 | 0.368 | 0.838 |
| Per SD decrease ${ }^{\text {c }}$ | 1.07 (0.95, 1.20) | 1.01 (0.89, 1.14) | 0.85 (0.69, 1.04) | 0.80 (0.64, 0.99) | 1.23 (1.00, 1.52) | 1.24 (0.99, 1.55) | 1.16 (0.94, 1.43) | 1.04 (0.83, 1.30) |

${ }^{\text {a }}$ Base model is adjusted for age, clinic, and season of blood draw.
${ }^{b}$ Multivariate model is adjusted for age, clinic, season of blood draw, serum calcium and phosphate, GFR, percentage body fat, weight, race, health status, presence of at least one medical condition, alcohol use, education, activity level (PASE score), marital status, and presence of a functional or mobility limitation.
${ }^{\text {c }}$ The sD for $25(\mathrm{OH}) \mathrm{D}$ is $7.98 \mathrm{ng} / \mathrm{ml}$.


FIG. 1. Adjusted survival plots for participants in the MrOS study, by quartile of total intact PTH or $25(\mathrm{OH})$ D. Plots are adjusted for age, clinic, season of blood draw, serum calcium and phosphate, GFR, percentage body fat, weight, race, health status, presence of at least one medical condition, alcohol use, education, activity level (PASE score), marital status, and presence of a functional or mobility limitation.
icant. How these somewhat small differences influenced mortality risk is unclear; multivariate adjustment removes the confounding effects of these imbalances.

The relation between PTH and $25(\mathrm{OH}) \mathrm{D}$ levels is complex, and most studies, including MrOS, show a modest association between these hormones. In our data, inclusion or exclusion of $25(\mathrm{OH}) \mathrm{D}$ levels in models estimating the association between PTH and mortality did not alter the effect estimate (and vice versa), suggesting that $25(\mathrm{OH}) \mathrm{D}$ and PTH may have statistically independent effects despite the biological link between them. Additionally, we did not find evidence for a statistical interaction between $25(\mathrm{OH}) \mathrm{D}$ and PTH, indicating that the association between PTH and mortality is consistent across all levels of $25(\mathrm{OH}) \mathrm{D}$ (and vice versa).

The finding that low $25(\mathrm{OH}) \mathrm{D}$ levels are not associated with increased risk of overall mortality contrasts with a meta-analysis of randomized trials (12) and large, popu-lation-based studies, such as the National Health and Nutrition Examination Survey (NHANES) (7) in which higher $25(\mathrm{OH})$ D levels were related to a lower risk of mortality. Several subtle differences between the current study and previous reports may cumulatively explain the divergent results in MrOS. First, in the MrOS study, $25(\mathrm{OH}) \mathrm{D}$ levels were assessed using liquid chromatography/mass spectroscopy, whereas most other studies have used radioimmunological or competitive binding assays. Second, most other observational studies had less well-characterized participants, which may have limited the ability to assess the role of confounding. For example, few studies have measures of adiposity (such as percentage body fat) aside from BMI. BMI is not a perfect measure of adiposity status $(31,32)$, and vitamin D status is associated with adiposity, so studies with better assessment of adiposity might be able to better control for the confounding influence of this factor. However, in MrOS, even the minimally adjusted association between $25(\mathrm{OH})$ D and mortality was not suggestive of an association, so the lack of adjustment for potential confounders in other studies is not likely to explain the difference in the findings. Also, although one quarter of the MrOS men are vitamin D deficient [25(OH)D $<20 \mathrm{ng} / \mathrm{ml}]$, very few men ( $<5 \%$ ) had $25(\mathrm{OH}) \mathrm{D}$ levels below $10 \mathrm{ng} / \mathrm{ml}$. If the association between vitamin D status and mortality is driven by individuals in this extremely low range, we may not have
TABLE 4. HRs ( $95 \% \mathrm{Cls}$ ) for mortality by categories of total intact PTH in the MrOS study

|  | All-cause mortality |  | Cancer |  | Cardiovascular |  | Noncancer, noncardiovascular |  |
| :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: | :---: |
|  | Base model ${ }^{\text {a }}$ | Multivariate ${ }^{\text {b }}$ | Base model ${ }^{\text {a }}$ | Multivariate ${ }^{\text {b }}$ | Base model ${ }^{\text {a }}$ | Multivariate ${ }^{\text {b }}$ | Base model ${ }^{\text {a }}$ | Multivariate ${ }^{\text {b }}$ |
| Quartil |  |  |  |  |  |  |  |  |
| $\begin{aligned} & \mathrm{Q} 1,<23.6 \\ & \mathrm{pg} / \mathrm{ml}(\mathrm{n}=372) \end{aligned}$ | 1.00 (referent) | 1.00 (referent) | 1.00 (referent) | 1.00 (referent) | 1.00 (referent) | 1.00 (referent) | 1.00 (referent) | 1.00 (referent) |
| $\begin{aligned} & \text { Q2, } 23.6 \text { to } \\ & <29.5 \mathrm{pg} / \mathrm{ml} \\ & (\mathrm{n}=373) \end{aligned}$ | 0.96 (0.69, 1.36) | 1.03 (0.73, 1.45) | 0.94 (0.51, 1.73) | 1.04 (0.56, 1.93) | 1.35 (0.72, 2.52) | 1.42 (0.75, 2.67) | 0.66 (0.36, 1.23) | 0.75 (0.39, 1.41) |
| $\begin{aligned} & \text { Q3, } 29.5 \text { to } \\ & <38.5 \mathrm{pg} / \mathrm{ml} \\ & (\mathrm{n}=372) \end{aligned}$ | 1.33 (0.97, 1.82) | 1.39 (1.00, 1.92) | 1.75 (1.03, 2.98) | 1.95 (1.12, 3.41) | 1.37 (0.74, 2.52) | 1.30 (0.69, 2.43) | 0.91 (0.52, 1.60) | 1.01 (0.56, 1.82) |
| $\begin{aligned} & \mathrm{Q} 4, \geq 38.5 \\ & \mathrm{pg} / \mathrm{ml}(\mathrm{n}=373) \end{aligned}$ | 1.34 (0.99, 1.82) | 1.32 (0.95, 1.84) | 0.77 (0.41, 1.45) | 0.94 (0.48, 1.83) | 2.04 (1.17, 3.55) | 1.51 (0.83, 2.76) | 1.25 (0.76, 2.06) | 1.33 (0.77, 2.31) |
| $P$ for trend | 0.019 | 0.039 | 0.955 | 0.499 | 0.011 | 0.235 | 0.235 | 0.212 |
| $\begin{aligned} & \text { Clinical categories } \\ & <66 \mathrm{pg} / \mathrm{ml} \\ & (\mathrm{n}=51) \end{aligned}$ | 1.00 (referent) | 1.00 (referent) | 1.00 (referent) | 1.00 (referent) | 1.00 (referent) | 1.00 (referent) | 1.00 (referent) | 1.00 (referent) |
| $\begin{aligned} & \geq 66 \mathrm{pg} / \mathrm{ml} \\ & (\mathrm{n}=1439) \end{aligned}$ | 2.15 (1.42, 3.24) | 1.64 (1.06, 2.56) | 0.59 (0.15, 2.41) | 0.62 (0.15, 2.59) | 3.92 (2.26, 6.80) | 1.82 (0.96, 3.43) | 1.86 (0.86, 4.03) | 1.76 (0.78, 4.00) |
| Per sD increase in log-transformed PTH ${ }^{\text {c }}$ | 1.19 (1.07, 1.32) | 1.15 (1.03, 1.29) | 0.97 (0.79, 1.19) | 1.03 (0.82, 1.29) | 1.46 (1.25, 1.71) | 1.21 (1.00, 1.45) | 1.12 (0.92, 1.35) | 1.15 (1.03, 1.29) |

[^0]had the ability to detect an effect due to the limited number of men with very low $25(\mathrm{OH}) \mathrm{D}$ levels.

Unexpectedly, there was a suggestion of a protective association between low levels of $25(\mathrm{OH}) \mathrm{D}$ and decreased risk of cancer mortality, a result that is in the opposite direction of the many studies in the literature and our initial hypothesis. Most studies have reported either no association between $25(\mathrm{OH}) \mathrm{D}$ levels and cancer mortality $(7,10,13,33)$ or increased cancer incidence or mortality for individuals with lower $25(\mathrm{OH}) \mathrm{D}$ levels $(34,35)$. However, a few reports have suggested that lower levels of $25(\mathrm{OH}) \mathrm{D}$ are associated with decreased cancer risk, such as a study of pancreatic cancer in Finish smokers (36). Only two randomized studies have examined the association between vitamin $D$ supplementation and cancer incidence with mixed results. The Women's Health Initiative found no association between vitamin D plus calcium and incidence of colorectal cancer $v$ s. placebo (37); a study of Nebraskan postmenopausal women found that both vitamin D plus calcium and calcium alone reduced overall cancer incidence $v$ s. placebo (38). Additionally, metaanalyses of observational studies suggest that vitamin D status may have different relationships with different cancer types; for example, one meta-analysis found support for an inverse association between $25(\mathrm{OH}) \mathrm{D}$ levels and colorectal cancer (39), whereas another report by the same authors found no association between $25(\mathrm{OH}) \mathrm{D}$ levels and prostate cancer (40). Finally, the association we observed was only marginally significant, and we made multiple comparisons, so the cancer mortality results should be interpreted with caution. At a minimum, our results provide no support for a role of increased $25(\mathrm{OH}) \mathrm{D}$ levels in cancer mortality prevention. However, the only way to unequivocally establish a causal association between vitamin D status and cancer outcomes is through a randomized trial.

This study had a number of strengths. Both $25(\mathrm{OH}) \mathrm{D}$ and PTH were measured using state-of-the-art assays in a large cohort of well-characterized, community-dwelling older men with nearly complete follow-up. However, a number of limitations must be noted. First, ascertainment of cause of death, completed centrally, used death certificates and discharge summaries that may have led to misclassification of the cause of death for some participants. Second, the number of participants who died from each of the major causes examined in this analysis is fairly small, resulting in HRs with wide CIs. Third, the small number of cancer and cardiovascular events precluded a meaningful examination of the association between $25(\mathrm{OH}) \mathrm{D}$ and PTH and subtypes of cancer or cardiovascular mortality. This is a particularly important limitation given that vi$\operatorname{tamin} \mathrm{D}$ status may have different associations with dif-
ferent cancer types. Additionally, we cannot exclude the possibility that some of the men with the very highest PTH levels had hyperparathyroidism. Finally, this study was conducted in relatively healthy men, so generalization of these results to other populations may be limited.

In summary, and in contrast to many but not all previous reports, lower $25(\mathrm{OH}) \mathrm{D}$ levels were not associated with an increased risk of all-cause or cause-specific mortality in community-dwelling older men. We found that higher PTH levels were associated with a modest increase in mortality risk, likely due to increased cardiovascular death. Randomized trials are needed to determine the causal relation between vitamin D and mortality.

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[^0]:    ${ }^{\text {a }}$ Base model is adjusted for age, clinic, and season of blood draw.
    ${ }^{b}$ Multivariate model is adjusted for age, clinic, season of blood draw, serum calcium and phosphate, GFR, percentage body fat, weight, race, health status, presence of at least one medical condition, alcohol use, education, activity level (PASE score), marital status, and presence of a functional or mobility limitation.
    ${ }^{\text {c }}$ sD for $\log$-transformed PTH is $0.394 \log (\mathrm{pg} / \mathrm{ml})$.

